

Financial Time Series – Nonparametric methods in time series

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- Highly data dependent, risk of overfitting
- No/few parameters/models \implies no error distribution

General time series model

- Let $X = (X_t, t \in \mathbb{Z})$ be given by

$$X_t = m(X_{t-r}) + Z_t,$$

where $Z \sim \text{IID}(0, \sigma^2)$ and m is an arbitrary, smooth, but unknown function

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- For Y independent of Z , $m(X_t) \approx \mathbb{E}(Y_t|X_t)$
- When $X = x$ is constant and independent of Z ,

$$y_t = m(x) + Z_t$$

and taking the sample average yields

$$n^{-1} \sum_{t=1}^n y_t = m(x) + n^{-1} \sum_{t=1}^n Z_t.$$

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- Otherwise: use a weighted average of y ,

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where $w_t(x)$ are larger for y_t with x_t close to x and smaller otherwise and $\sum_{t=1}^n w_t(x) = 1$.

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- *Kernel* $K : \mathbb{R} \rightarrow \mathbb{R}^+$. Typically a density function, $\int K(z) \, dz = 1$.
- Rescale by *bandwidth*:

$$K_h(x) = h^{-1} K(xh^{-1}).$$

Still:

$$\int K_h(z) \, dz = 1.$$

Kernel regression

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Example: The *Gaussian kernel*

$$K_h(x) := (2\pi h^2)^{-1/2} \exp(-(2h^2)^{-1} x^2)$$

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$$K_h(x) := 0.75 h^{-1} (1 - (x/h)^2) I(|x/h| \leq 1),$$

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For the latter: $\hat{m}(x_t) \rightarrow y_t$ for $h \rightarrow 0$ and $\hat{m}(x_t) \rightarrow \bar{y}$ for $h \rightarrow \infty$

Bandwidth selection

Method (Bandwidth selection with MISE)

Minimize the *mean integrated squared error*:

$$\text{MISE} := \mathbb{E} \left(\int_{-\infty}^{\infty} (\hat{m}(x) - m(x))^2 \, dx \right),$$

where m is the true function and \hat{m} the estimator which depends on h .

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Expand this to derive an optimal bandwidth that depends on unknown quantities.

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where m is the true function and \hat{m} the estimator which depends on h . Expand this to derive an optimal bandwidth that depends on unknown quantities. Estimate these by preliminary smoothing, i.e., computing \hat{m} with a reference bandwidth selector. Common choices:

$$\hat{h}_{\text{opt}} = \begin{cases} 1.06 s n^{-1/5} & \text{for the Gaussian kernel,} \\ 2.34 s n^{-1/5} & \text{for the Epanechnikov kernel,} \end{cases}$$

where s is the sample standard error of $(x_t)_{t=1}^n$.

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Method (Bandwidth selection with cross validation)

Leave-one-out cross validation omits (x_j, y_j) . The other data points are used to find:

$$\hat{m}_{h,j}(x_j) := \sum_{t \neq j} w_t(x_j) y_t \approx y_j.$$

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Repeat for the remaining $n - 1$ points and set

$$\text{CV}(h) := \sum_{j=1}^n (y_j - \hat{m}_{h,j}(x_j))^2 W(x_j),$$

where W is a nonnegative weight function satisfying

$$\sum_{j=1}^n W(x_j) = 1.$$

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where W is a nonnegative weight function satisfying

$\sum_{j=1}^n W(x_j) = 1$. Set $h_{\text{opt}} := \arg \min_h \text{CV}(h)$.

Local linear regression

- Idea: $\hat{m}(x)$ can equally be defined as

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- The *local linear regression method*:

$$(\hat{a}, \hat{b}) = \arg \min_{a,b} \sum_{t=1}^n (y_t - a - b(x - x_t))^2 K_h(x - x_t)$$

- Beats kernel regression if m is twice continuously differentiable

Local linear regression

A closed form solution:

$$\hat{a} = \frac{\sum_{t=1}^n w_t(x) y_t}{\sum_{t=1}^n w_t(x)} \text{ and } \hat{b} = \frac{\sum_{t=1}^n \tilde{w}_t(x) y_t}{\sum_{t=1}^n w_t(x)}$$

where

$$w_t(x) := K_h(x - x_t)(s_{n,2}(x) - (x - x_t)s_{n,1}(x)),$$

$$\tilde{w}_t(x) := K_h(x - x_t)((x - x_t)s_{n,0}(x) - s_{n,1}(x)),$$

and

$$s_{n,j}(x) := \sum_{t=1}^n K_h(x - x_t)(x - x_t)^j$$

for $j = 0, 1, 2$.

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Extend to more complex autoregressive models like

$$X_t = m_1(X_{t-1}) + m_2(X_{t-2}) + \dots + m_k(X_{t-k}) + Z_t,$$

where $(m_i)_{i=1}^k$ is a sequence of smooth functions, or

$$X_t = m(X_{t-1}, \dots, X_{t-k}) + Z_t,$$

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Multivariate kernels $K_h : \mathbb{R}^k \rightarrow \mathbb{R}$, e.g., multivariate Gaussian density:

$$K_h(\mathbf{x}) = (2\pi h^2)^{-k/2} (\det \Sigma)^{-1/2} \exp(-(2h^2)^{-1} \mathbf{x}' \Sigma^{-1} \mathbf{x}).$$