

Lecture_22

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Options and Mathematics: Lecture 22

December 9, 2020

Exercises

Exercise 6.6

Let $\{W(t)\}_{t \geq 0}$ be a Brownian motion. Show that

$$\text{Cov}[W(s), W(t)] = \min(s, t), \quad \text{for all } s, t \geq 0.$$

(Solution can be found in the book)

SOLUTION: RECALL THAT $\text{Cov}(X, Y) = E[XY] - E[X]E[Y]$

ASSUME $s \geq t$ (THE SOLUTION IS SIMILAR FOR $s \leq t$)

$$\begin{aligned} \text{Cov}[W(s), W(t)] &= E[W(s)W(t)] - E[W(s)]E[W(t)] = \\ &\quad \downarrow \text{SINCE } W(t) \in N(0, t) \text{ THEN } E[W(s)] = E[W(t)] = 0 \\ &= E[W(s)W(t)] = E[(\underbrace{W(s)}_{0} - \underbrace{W(t)}_{0}) \underbrace{W(t)}_{0}] + E[W(t)^2] \\ &= E[(W(s) - W(t))(W(t) - \underbrace{W(0)}_{0})] + (E[W(t)^2] - \underbrace{(E[W(t)])^2}_{0}) \end{aligned}$$

$$= E[\underbrace{W(s)}_0 \underbrace{W(t)}_0] + V_{\text{a.s.}}[W(t)] = t = \min(s, t)$$

WE ASSUMED $s \geq t$

$$= E[W(s)/W(t)] E[X(s)] + V_{A \in [W(t)]} = t = \min(s, t)$$

WE ASSUMED $s > t$

Exercise 6.7

Let $\{W(t)\}_{t \geq 0}$ be a \mathbb{P} -Brownian motion and $T > 0$. Given a differentiable function $\theta : (0, \infty) \rightarrow \mathbb{R}$, define

$$Z_\theta = \exp \left(-\theta(T)W(T) + \int_0^T \theta'(s)W(s) ds - \frac{1}{2} \int_0^T \theta^2(s) ds \right).$$

Show that $\mathbb{P}_\theta(A) = \mathbb{E}[Z_\theta \mathbb{I}_A]$ is a probability measure equivalent to \mathbb{P} .

HINT: You need Theorem 6.6:

Let $g : (0, \infty) \rightarrow \mathbb{R}$ be a differentiable function and let

$$X(t) = g(t)W(t) - \int_0^t g'(s)W(s) ds.$$

(REPLACE $g(t) = \theta(t)$ AND $t = T$)

Then

$$X(t) \in \mathcal{N}(0, \Delta(t)), \quad \Delta(t) = \int_0^t g(s)^2 ds.$$

IF WE TAKE
 $\theta(t) = \theta$ CONST.
 THEN WE GO
 BACK TO
 THIS CASE

SOLUTION: RECALL THAT IF $Z_\theta = e^{-\theta W(T) - \frac{1}{2} \theta^2 T}$, $\theta \in \mathbb{R}$,
 THEN $\mathbb{P}_\theta(A) = \mathbb{E}[Z_\theta \mathbb{I}_A]$ IS, BY RADON-NIKODYM THEOREM,
 A PROBABILITY EQUIVALENT TO \mathbb{P} .
 (GIRSANOV'S PROBABILITY)

TO SOLVE THIS EXERCISE WE HAVE TO SHOW THAT ALSO THE NEW Z_θ SATISFIES $Z_\theta > 0$, $\mathbb{E}[Z_\theta] = 1$.

CLEARLY $Z_\theta > 0$; MOREOVER:

$$\mathbb{E}[Z_\theta] = e^{-\frac{1}{2} \int_0^T \theta(s)^2 ds} \mathbb{E} \left[e^{-\theta(T)W(T) + \int_0^T \theta'(s)W(s) ds} \right]$$

$X(T) \in \mathcal{N}(0, \Delta(T))$

$$\text{WHERE } \Delta(T) = \int_0^T \theta(s)^2 ds. \quad \text{HENCE}$$

$$\mathbb{E}[Z_\theta] = e^{-\frac{1}{2} \int_0^T \theta(s)^2 ds} \int_{-\infty}^{\infty} e^{-x - \frac{x^2}{2 \Delta(T)}} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx = e^{-\frac{1}{2} \int_0^T \theta(s)^2 ds} \int_{-\infty}^{\infty} e^{-\frac{x^2}{2 \Delta(T)}} \frac{dx}{\sqrt{2\pi}}$$

$$\begin{aligned}
&= e^{-\frac{1}{2} \int_0^T (\theta(s))^2 ds} \int_{\mathbb{R}} e^{-\frac{1}{2\Delta(\tau)}(x^2 + 2x\Delta(\tau) + \underline{\Delta(\tau)^2}) + \frac{\Delta(\tau)}{2}} \frac{dx}{\sqrt{2\pi}} \\
&= e^{-\frac{1}{2}\Delta(\tau)} e^{\frac{1}{2}\Delta(\tau)} \int_{\mathbb{R}} e^{-\frac{1}{2\Delta(\tau)}(x + \Delta(\tau))^2} \frac{dx}{\sqrt{2\pi}} = 1
\end{aligned}$$

\downarrow Density of a RV in $N(-\Delta(\tau), \Delta(\tau))$

Exercise 6.9

Suppose that at time $t = 0$ it is assumed that the stock price is described by a geometric Brownian motion

$$S(t) = S_0 e^{(r - \frac{\sigma^2}{2})t + \sigma W(t)}$$

\uparrow
THERE IS
NO σ HERE!

$$\text{IN } P_0: S(t) = S(0)e^{rt + \sigma W(t)}$$

THE VALUE OF OPTIONS
ON THE STOCK DOES
NOT DEPEND ON d . HOWEVER
IT DEPENDS
ON σ !!

in the interval $[0, T]$. Given any arbitrary subinterval $[t_0, t] \subset [0, T]$ with length $\tau = t - t_0$, define the random variable

$$\sigma_\tau^2(t) = \frac{1}{h(n-1)} \sum_{i=1}^n (\hat{R}_i - \bar{R})^2,$$

where $t_0 < t_1 < t_2 < \dots < t_n = t$ is a partition of $[t_0, t]$ with $h = t_i - t_{i-1}$ and \hat{R}_i, \bar{R} are given by

$$\begin{aligned}
\text{RANDOM VARIABLES} \left\{ \begin{array}{l} \hat{R}_i = \log S(t_i) - \log S(t_{i-1}) = \log \left(\frac{S(t_i)}{S(t_{i-1})} \right), \quad i = 1, \dots, n. \\ \bar{R}(t) = \frac{1}{n} \sum_{i=1}^n \hat{R}_i = \frac{1}{n} \log \left(\frac{S(t)}{S(t_0)} \right). \end{array} \right. \rightarrow \text{LOG-RETURN} \\ \text{ON } [t_{i-1}, t_i] \rightarrow \text{MEAN OF LOG RETURN}
\end{aligned}$$

Show that

$$\rightarrow \mathbb{E}[\sigma_\tau^2(t)] = \sigma^2$$

In other words, σ^2 is the expected value of the τ -historical variance at any time $t \in [0, T]$.

(Solution can be found in the book)

$$C(T, x, K, \tau) = x \Phi(d_{(+)}) - K e^{-r\tau} \Phi(d_{(-)})$$

DO IT YOURSELF

$$d_{(\pm)} = \frac{\log \frac{x}{K} + (\tau + \frac{1}{2}\sigma^2)\pm}{\sigma \sqrt{\tau}}$$

$$\tau = T - t$$

Let $C(t, x, K, T)$ be the Black-Scholes pricing function of the call with strike K and maturity T .

Prove that

$$\lim_{\sigma \rightarrow 0^+} C(t, x, K, T) = (x - Ke^{-rt})_+, \quad \lim_{\sigma \rightarrow \infty} C(t, x, K, T) = x.$$

Compute also the following limits:

$$\lim_{K \rightarrow 0^+} C(t, x, K, T), \quad \lim_{K \rightarrow +\infty} C(t, x, K, T), \quad \lim_{\tau \rightarrow +\infty} C(t, x, K, T), \quad \lim_{x \rightarrow 0^+} C(t, x, K, T)$$

and show that $C(t, x, K, T)$ is asymptotic to $x - Ke^{-rt}$ as $x \rightarrow \infty$. Compute the same limits for put options.

 (Solution can be found in the book)

HINT: REMEMBER THAT $\Phi(z) \rightarrow 1$
 AS $z \rightarrow \infty$
 AND $\Phi(z) \rightarrow 0$
 $z \rightarrow -\infty$

Exercise 6.13

A **binary** (or **digital**) call option with strike K and maturity T pays-off the buyer if and only if $S(T) > K$. If the pay-off is a fixed amount of cash L , then the binary call option is said to be “cash-settled”, while if the pay-off is the stock itself then the option is said to be “physically settled”. Compute the Black-Scholes price of the cash-settled binary call option and the number of shares on the stock in the self-financing hedging portfolio.

(Solution can be found in the book)

$$\text{SOLUTION: } Y = L H(S(T) - K) = g(S(T))$$

$$g(z) = L H(z - K)$$

$$P_{Y(t)} = N(+, S(t)), \quad N(+x) = e^{-rt} \int_R^{\infty} g(xe^{(r-\frac{\sigma^2}{2})t + rt^2/2}) e^{\frac{-\frac{1}{2}\gamma^2}{\sigma^2 t}} dy$$

$$= e^{-rt} \int_R^{\infty} L H(xe^{(r-\frac{\sigma^2}{2})t + rt^2/2} - K) e^{-\frac{1}{2}\gamma^2 \frac{dy}{\sigma^2 t}} \quad \gamma = T - t$$

$$\left\{ x e^{(r-\frac{\sigma^2}{2})t + rt^2/2} - K > 0 \text{ IFF } \gamma > \frac{\log \frac{x}{K} - (r - \frac{\sigma^2}{2})t}{\frac{1}{2}\gamma^2} = -d_1(-) \right\}$$

$$= e^{-rt} \int_{-d_1(-)}^{\infty} L e^{-\frac{1}{2}\gamma^2 \frac{dy}{\sigma^2 t}} \quad 5 = L e^{-rt} \int_{-\infty}^{d_1(-)} e^{-\frac{1}{2}\gamma^2 \frac{dy}{\sigma^2 t}}$$

$$d_1(-) = \frac{\log \frac{x}{K} + (r - \frac{\sigma^2}{2})t}{\sigma \sqrt{t}} = \frac{\log \frac{x}{K}}{\sigma \sqrt{t}} + \underbrace{C}_{\text{DOESN'T DEPEND ON X}} = L e^{-rt} \bar{\Phi}(d_1(-))$$

$$h_S(t) = \Delta H(SH) \quad , \quad \Delta(t,x) = \partial_x \Phi(t,x) \quad \text{and} \\ \Delta(H_t, x) = \partial_x [L e^{-rt} \Phi(d_{t-})] = L e^{-rt} \Phi'(d_{t-}) \partial_x d_{t-} \\ = L e^{-rt} \varphi(d_{t-}) \left(\frac{1}{x \sigma \sqrt{r}} \right) > 0 \text{ where } \varphi(z) = \frac{e^{-\frac{1}{2}z^2}}{\sqrt{2\pi}}$$

Exercise 6.15

Consider the European derivative with maturity T and pay-off Y given by

$$Y = k + S(T) \log S(T),$$

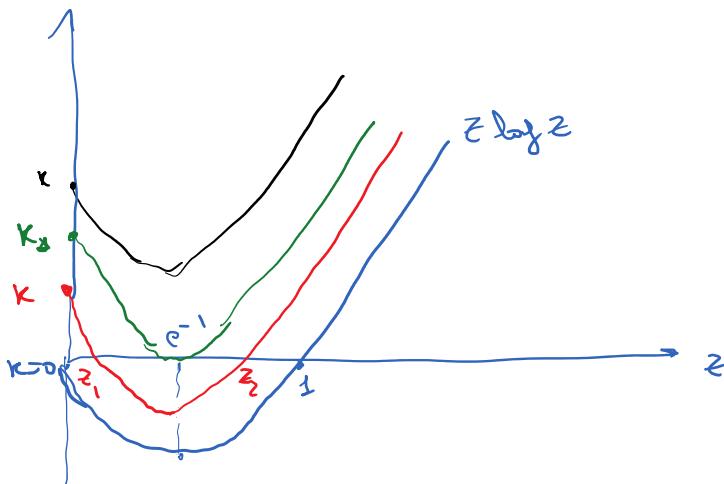
$$= \Phi'(z)$$

where $k > 0$ is a constant. Find the Black-Scholes price of the derivative at time $t < T$ and the self-financing hedging portfolio. Find the probability that the derivative expires in the money.

(Solution can be found in the book) (DO IT YOURSELF)

HINT: $Y = g(S(T))$

$$g(z) = k + z \log z$$



$$g'(z) = \log z + 1 \\ \Rightarrow z = e^{-1}$$

THERE IS
A VALUE OF k ,
SAY $k = k_*$,
SUCH THAT
IF $k > k_*$, THEN

$$Y > 0, \text{ AND SO} \\ P(Y > 0) = 1$$

FOR $0 < k < k_*$ THERE EXIST
 $z_1 < z_2$ SUCH THAT $g(z) > 0$
IFF $z < z_1$ OR $z > z_2$, HENCE 6

$$Y > 0 \Leftrightarrow S(T) < z_1 \text{ OR } S(T) > z_2$$

$$\text{HENCE } P(Y > 0) = P(S(T) < z_1) + P(S(T) > z_2)$$

Solution 6.17

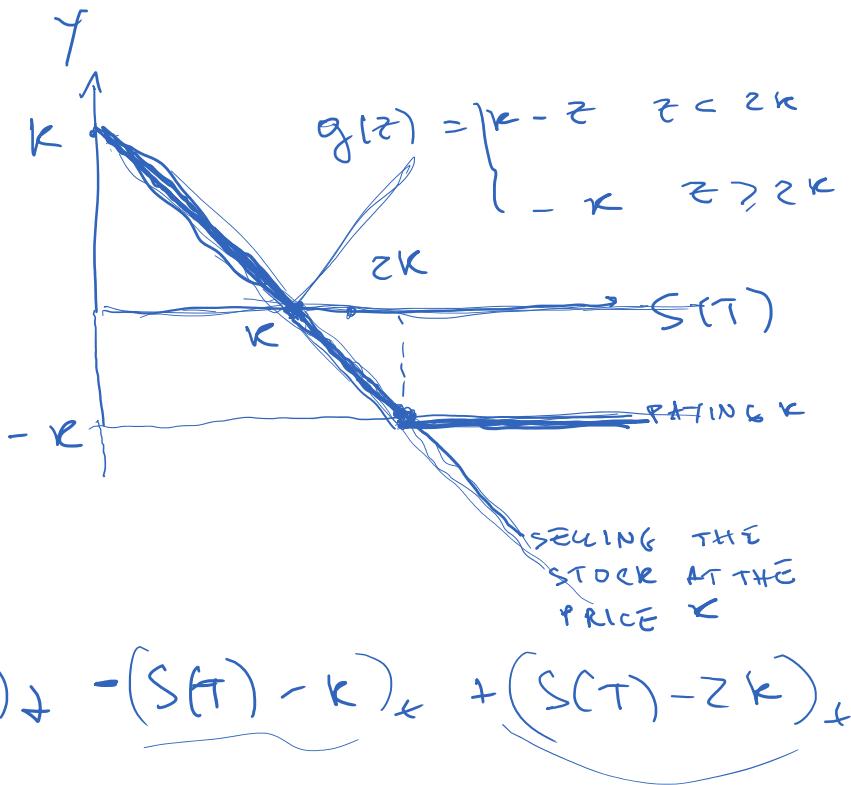
Let $K > 0$. A European style derivative on a stock with maturity $T > 0$ gives to its owner the right to choose between selling the stock for the price K at time T or paying the amount K at time T . Draw the pay-off function of the derivative. Compute the Black-Scholes price of the derivative. Show that there exists a value K_* of K such that the Black-Scholes price of the derivative is zero.

READ THE
SOLUTION
IN THE
BOOK.

(Solution can be found in the book)

SOLUTION
FIRST DRAW
THE GRAPH OF
THE PAY-OFF

$$\gamma = g(S(T))$$



$$\gamma = (\underbrace{K - S(T)}_{+}) - \underbrace{(S(T) - K)}_{\times} + \underbrace{(S(T) - zK)}_{+}$$

$$\Rightarrow \Pi_Y(t) = \underbrace{\mathbb{P}(t, S(t), K, T) - C(t, S(t), K, T)}_{+} + C(t, S(t), zK, T) = \underbrace{K e^{-rT} - S(t)}_{\text{PUT-CALL PARITY}} + C(t, S(t), zK, T)$$

$$\therefore \Pi_Y(t) = K e^{-rT} - S(t) + C(t, S(t), zK, T)$$

Solution to Exercise 6.9

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$$S(t_i) = S_0 e^{\alpha t_i + \sigma W(t_i)}$$

$$\log S(t_i) - \log S(t_{i-1}) = \alpha(t_i - t_{i-1}) + \sigma(W(t_i) - W(t_{i-1}))$$

$$\hat{R}_i^{(1)} =$$

$$R = \frac{1}{n} \sum \hat{R}_i = \frac{1}{n} \log \frac{S(t)}{S(t_0)} = \frac{1}{n} (\alpha(t - t_0) + \sigma(W(t) - W(t_0)))$$

$$\hat{\sigma}_x^2(t) = \frac{1}{n(n-1)} \sum_{i=1}^n \left[\alpha(t_i - t_{i-1}) + \sigma(W(t_i) - W(t_{i-1})) - \frac{1}{n} (\alpha(t - t_0) + \sigma(W(t) - W(t_0))) \right]^2$$

We want to show that $E[\hat{\sigma}_x^2(t)] = \sigma^2$

BREAK until 2.13 PM.

$$E[\hat{\sigma}_x^2(t)] = \frac{1}{n(n-1)} \sigma^2 E \left[\left(\sum_{i=1}^n (W(t_i) - W(t_{i-1}))^2 + \frac{1}{n} (W(t) - W(t_0))^2 \right) - \frac{2}{n} (W(t) - W(t_0)) (W(t) - W(t_0)) \right]$$

$$= \frac{1}{n(n-1)} \sigma^2 E \left[\left(\sum_{i=1}^n (W(t_i) - W(t_{i-1}))^2 \right) + \frac{1}{n} (W(t) - W(t_0))^2 - \frac{2}{n} (W(t) - W(t_0)) \sum_{i=1}^n (W(t_i) - W(t_{i-1})) \right]$$

THIS SUM TELESCOPES,

$$= \frac{1}{n(n-1)} \sigma^2 E \left[\left(\sum_{i=1}^n (W(t_i) - W(t_{i-1}))^2 \right) - \frac{1}{n} (W(t) - W(t_0))^2 \right]$$

$$= \frac{1}{n(n-1)} \sigma^2 \left(\sum_{i=1}^n E[(W(t_i) - W(t_{i-1}))^2] - \frac{1}{n} E[(W(t) - W(t_0))^2] \right)$$

$$\begin{aligned}
 & \text{VAR} [w(t_i) - w(t_{i-1})] \quad \text{VAR} [w(t) - w(t_0)] \\
 &= t_i - t_{i-1} = h \quad = (t - t_0) \\
 &= \frac{1}{n(n-1)} \sigma^2 \left(h n - \underbrace{\frac{1}{n}(t - t_0)}_h \right) = \frac{1}{n} \frac{\sigma^2}{(n-1)} n(n-1) = \sigma^2
 \end{aligned}$$

