

Lecture_23

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Options and Mathematics: Lecture 23

December 10, 2020

The Asian option

For European call and put options, and for other simple standard European derivatives, the Black-Scholes pricing formula can be reduced to a simple expression in terms of the standard normal distribution.

For non-standard derivatives, i.e., when the pay-off depends on the price of the stock at different times (and not only at maturity), this reduction is in general not possible.

Nevertheless the risk-neutral pricing formula can be used to compute numerically the Black-Scholes price of non-standard derivatives using the so called **Monte Carlo method**.

We illustrate the procedure in the important case of the Asian option.

Recall that the Asian call, resp. put, option in the time-continuum case is defined as the non-standard European derivative with pay-off

NON-STANDARD
DERIVATIVES

$$Y_{\text{call}} = \left(\frac{1}{T} \int_0^T S(t) dt - K \right)_+, \quad \text{resp.} \quad Y_{\text{put}} = \left(K - \frac{1}{T} \int_0^T S(t) dt \right)_+,$$

CASH-
SETTLED
OTC OPTIONS

ARITHMETIC AVERAGE OF $S(t)$ IN THE INTERVAL $[0, T]$
where $K > 0$ is the strike price of the option.

THE BLACK-SCHOLES PRICE AT TIME $t=0$ IS

$$\pi_{AC}(0) = e^{-rT} \mathbb{E}_Q[Y_{\text{call}}] \quad \pi_{AP}(0) = e^{-rT} \mathbb{E}_Q[Y_{\text{put}}]$$

FOR THE STANDARD CALL THE PAY-OFF IS

$$\underline{(S(T) - K)_+}$$

The Black-Scholes price at time $t = 0$ of these options are given respectively by

$$\Pi_{AC}(0) = e^{-rT} \mathbb{E}_q[Y_{\text{call}}], \quad \Pi_{AP}(0) = e^{-rT} \mathbb{E}_q[Y_{\text{call}}].$$

Exercise 6.23

Derive the following put-call parity identity:

$$\rightarrow \Pi_{AC}(0) - \Pi_{AP}(0) = e^{-rT} \left(\frac{e^{rT} - 1}{rT} S_0 - K \right).$$

$$\frac{\sum_{i=1}^N x_i}{N} = \text{ARITHMETIC AVERAGE OF } x_1, \dots, x_N$$

$$\left(\prod_{i=1}^N x_i \right)^{1/N} = \text{GEOMETRIC AVERAGE OF } x_1, \dots, x_N$$

Exercise 6.24

The Asian call with geometric average is the non-standard European derivative with pay-off

$$Q = \left(e^{\frac{1}{T} \int_0^T \log S(t) dt} - K \right)_+.$$

→ ALSO EXIST IN REAL MARKETS (OTC)
→ THIS IS MORE SIMILAR TO A STANDARD CALL

Show that the Black-Scholes price at time $t = 0$ of this derivative is given by

$$\Pi_{AC}^{(G)}(0) = e^{-rT} (e^{qT} S_0 \Phi(d_1) - K \Phi(d_2))$$

where

$$q = \frac{1}{2} \left(r - \frac{\sigma^2}{6} \right) T, \quad d_2 = d_1 - \sigma \sqrt{\frac{T}{3}}, \quad d_1 = \frac{\log \frac{S_0}{K} + \frac{1}{2} \left(r + \frac{\sigma^2}{6} \right) T}{\sigma \sqrt{T/3}}.$$

Derive also the analogous formula the Black-Scholes price of the put option as well as the corresponding put-call parity.

HINT: You need Theorem 6.6 (and not 6.9 as written in the book!).

BLACK-SCHOLES PRICE AT $t=0$ OF THE EUROPEAN DERIVATIVE WITH PAY Y AND MATURITY $T > 0$ IS GIVEN BY THE RISK-NEUTRAL PRICING FORMULA $\pi_Y(0) = e^{-rT} \mathbb{E}_Q[Y]$ ←
 IF $Y = g(S(T))$, THIS PRICE CAN BE WRITTEN IN THE INTEGRAL FORM $\pi_Y(0) = \pi_0(S_0)$, $\pi_0(x) = e^{-rT} \int_{\mathbb{R}} g(xe^{(r-\frac{\sigma^2}{2})T + \sigma\sqrt{T}y}) e^{-\frac{1}{2}y^2} \frac{dy}{\sqrt{2\pi}}$

The Monte Carlo method

The Monte Carlo method is, in its simplest form, a numerical method to compute the expectation of a random variable.

Its mathematical validation is based on the **Law of Large Numbers**, which states the following: Suppose $\{X_i\}_{i \geq 1}$ is a sequence of i.i.d. random variables with expectation $\mathbb{E}[X_i] = \mu$. Then the sample average of the first n components of the sequence, i.e.,

i.i.d.
 INDEPENDENT
 IDENTICALLY
 DISTRIBUTED

$$\mathbb{P}(|\bar{X}_n - \mu| > \varepsilon) \rightarrow 0$$

$$\bar{X}_n = \frac{1}{n}(X_1 + X_2 + \dots + X_n),$$

FOR ALL $\varepsilon > 0$

converges (in probability) to μ as $n \rightarrow \infty$.

$$\bar{X}_n \rightarrow \mu \text{ as } n \rightarrow \infty$$

The law of large numbers can be used to justify the fact that if we are given a large number of independent trials X_1, \dots, X_n of the random variable X , then

$$\mathbb{E}[X] \approx \frac{1}{n}(X_1 + X_2 + \dots + X_n).$$

To measure how reliable is the approximation of $\mathbb{E}[X]$ given by the sample average, consider the standard deviation of the trials X_1, \dots, X_n :

$$s_X = \sqrt{\frac{1}{n-1} \sum_{i=1}^n (\bar{X} - X_i)^2}.$$

$$\bar{X} = \frac{X_1 + \dots + X_n}{n} = O(1)$$

$$O(1) = O(1) = O(1)$$

A simple application of the Central Limit Theorem proves that the random variable

$$F_{Y_n}(x) \xrightarrow{n \rightarrow \infty} \Phi(x)$$

$$Y_n = \frac{\mu - \bar{X}}{s_X / \sqrt{n}} \rightarrow Y \in N(0, 1)$$

$$\mathbb{E}[X] = \mu$$

converges in distribution to a standard normal random variable. We use this

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$$\text{IF } n \text{ IS SUFFICIENTLY LARGE } \mathbb{P}(a < Y_n < b) \approx \frac{1}{\sqrt{2\pi}} \int_a^b e^{-\frac{1}{2}x^2} dx$$

$\bar{X} \equiv$ MONTE-CARLO APPROXIMATION OF μ

$I \equiv 95\%$ CONFIDENCE INTERVAL

μ IS SOMEWHERE
HERE WITH
95% PROBABILITY

result to show that the true value μ of $\mathbb{E}[X]$ has about 95% probability to be in the interval

$$\bar{X} - 1.96 \frac{s}{\sqrt{n}} \quad \bar{X} + 1.96 \frac{s}{\sqrt{n}}$$

Indeed, for n large

$$\mu \in [\bar{X} - 1.96 \frac{s}{\sqrt{n}}, \bar{X} + 1.96 \frac{s}{\sqrt{n}}] \quad \text{WITH } 95\% \text{ PROBABILITY (IF } \mu \text{ IS LARGE ENOUGH)}$$

$$\mathbb{P} \left(-1.96 \leq \frac{\mu - \bar{X}}{s_X / \sqrt{n}} \leq 1.96 \right) \approx \int_{-1.96}^{1.96} e^{-x^2/2} \frac{dx}{\sqrt{2\pi}} \approx 0.95.$$

Application to the Asian option

Consider now a European derivative with pay-off Y at maturity T . We approximate the price at time $t = 0$ by

$$\Pi_Y(0) = e^{-rT} \mathbb{E}_Q[Y] \approx e^{-rT} \frac{Y_1 + \dots + Y_n}{n}$$

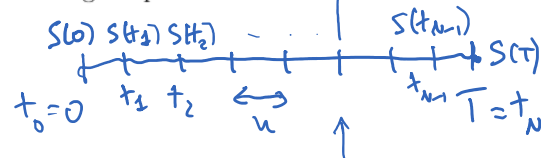
$$S(t) = \begin{cases} S(0)e^{rt + \sigma W(t)} & \text{in } \mathbb{P} \\ S(0)e^{(r - \frac{\sigma^2}{2})t + \sigma W^{(Q)}(t)} & \text{in } \mathbb{P}_Q \end{cases}$$

where Y_1, \dots, Y_n is a large number of independent pay-off trials.

As the pay-off depends on the path of the stock price, the trials Y_1, \dots, Y_n can be created by first generating a sample of paths for the stock price.

Letting $0 = t_0 < t_1 < \dots < t_N = T$ be a partition of the interval $[0, T]$ with size $t_i - t_{i-1} = h$, we may construct a sample of n paths of the geometric Brownian motion on the given partition with the following simple Matlab function:

```
function Path=StockPath(s,sigma,r,T,N,n)
h=T/N;
W=randn(n,N);
q=ones(n,N);
Path=s*exp((r-sigma^2/2)*h.*cumsum(q')+sigma*sqrt(h)*cumsum(W'));
Path=[s*ones(1,n);Path];
```



$$S(t_i) = S(0) e^{(r - \frac{\sigma^2}{2})t_i + \sigma W^{(Q)}(t_i)}$$

Note carefully that the stock price is modeled as a geometric Brownian motion with mean of log return $\alpha = r - \sigma^2/2$, which means that the geometric Brownian motion is defined in the risk-neutral probability.

This is of course correct, since the expectation that we want to compute is in the risk-neutral probability measure.

In the case of the Asian call option with strike K and maturity T the pay-off is given by

$$Y = \left(\frac{1}{T} \int_0^T S(t) dt - K \right)_+ \approx \left(\frac{1}{N} \sum_{i=1}^N S(t_i) - K \right)_+.$$

The following function computes the approximate price of the Asian option using the Monte Carlo method:

```
function [price, err]=MonteCarlo_AC(s,sigma,r,K,T,N,n)
tic
stockPath=StockPath(s,sigma,r,T,N,n);
payOff=max(0,mean(stockPath)-K);
price=exp(-r*T)*mean(payOff);
err=1.96*std(payOff)/sqrt(n);
toc
```

The function also returns the error in the 95% confidence interval, that is

$$\text{Err} = 1.96 \frac{s}{\sqrt{n}}$$

For example, by running the command

```
[price, err]=MonteCarlo_AC(100,0.5,0.05,100,1/2,100,1000000)
```

we get **price**=8.5799, **err**=0.0283, which means that the Black-Scholes price of the Asian option with the given parameters has 95% probability to be in the interval 8.5799 ± 0.0283 . The calculation took about 4 seconds. Note that the 95% confidence is $0.0565/8.5799 * 100 \approx 0.66\%$ of the price.

PRICE = $T \gamma(0)$
 $\text{Err} = 1.96 \frac{s}{\sqrt{n}}$

POINTS IN THE PARTITION
 ↳ # OF SAMPLE PATHS

2¹⁰⁰ POSSIBLE PATHS OF THE STOCK PRICE
 ↳ 10⁶

In order to reduce the error, i.e., to shrink the confidence interval, of the Monte Carlo approximation, one needs to either

- (i) increase the number of trials n or
- (ii) reduce the standard deviation s . Increasing n can be very costly in terms of computational time, hence the approach (ii) is more efficient.

There exist several methods to decrease the standard deviation of a Monte Carlo computation, which are known as variance reduction techniques. An example for the Asian option can be found in the book

BREAK UNTIL 25.15

Solution to Exercise 6.23

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$$\begin{aligned}\pi_{AC}(0) &= e^{-rT} \mathbb{E}_Q[Y_{CALL}] & Y_{CALL} &= \left(\frac{1}{T} \int_0^T S(t) dt - K \right)_+ \\ \pi_{AP}(0) &= e^{-rT} \mathbb{E}_Q[Y_{PUT}] & Y_{PUT} &= \left(K - \underbrace{\frac{1}{T} \int_0^T S(t) dt}_Z \right)_+\end{aligned}$$

$$\pi_{AC}(0) - \pi_{AP}(0) = e^{-rT} \mathbb{E}_Q[Y_{CALL} - Y_{PUT}]$$

$$= \left\{ (z - K)_+ - (K - z)_+ = z - K \text{ for all } z \right\}$$

$$= e^{-rT} \mathbb{E}_Q \left[\frac{1}{T} \int_0^T S(t) dt - K \right]$$

$$= e^{-rT} \mathbb{E}_Q \left[\frac{1}{T} \int_0^T S(t) dt \right] - K e^{-rT}$$

$$= e^{-rT} \frac{1}{T} \int_0^T \mathbb{E}_Q[S(t)] dt - K e^{-rT}$$

$$= e^{-rT} \frac{1}{T} \int_0^T \underbrace{\mathbb{E}_Q[e^{-rt} S(t)]}_{S^*(t)} e^{rt} dt - K e^{-rT}$$

$$(\mathbb{E}_Q[S^*(t)] = \mathbb{E}_Q[S^*(0)] = S(0)) \leftarrow$$

$$= e^{-rT} \frac{S_0}{T} \left(\int_0^T e^{rt} dt \right) - K e^{-rT}$$

$$= e^{-rT} \frac{S_0}{T} \frac{e^{rT} - 1}{r} - K e^{-rT}$$

$$= e^{-rT} \left[S_0 \frac{e^{rT} - 1}{rT} - K \right]$$

Solution to Exercise 6.24

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$$\Pi_{AC}^{(f)}(0) = e^{-rT} \mathbb{E}_Q[Q]$$

$$Q = \left(e^{\frac{1}{T} \int_0^T \log S(t) dt} - K \right)_+$$

$$S(t) = S_0 e^{rt + \sigma W(t)} \quad \text{in the PHYS. PROBABILITY } \mathbb{P}$$

$$= S_0 e^{(\underbrace{r - \frac{\sigma^2}{2}}_{\log S_0 + \log e^{(\dots)}})t + \sigma W^{(Q)}(t)} \quad \text{IN THE RISK-NEUTRAL (OR MARTINGALE) PROBABILITY } \mathbb{P}_Q$$

$$\log S(t) = \log S_0 + \left(r - \frac{\sigma^2}{2}\right)t + \sigma W^{(Q)}(t)$$

$$e^{\frac{1}{T} \int_0^T \log S(t) dt} = e^{\frac{1}{T} \int_0^T [\log S_0 + (r - \frac{\sigma^2}{2})t + \sigma W^{(Q)}(t)] dt} \quad \log S(t)$$

$$= e^{\left(\log S_0\right) + \left(r - \frac{\sigma^2}{2}\right)\frac{T}{2} + \frac{\sigma}{T} \int_0^T W^{(Q)}(t) dt}$$

$$= S_0 e^{\left(r - \frac{\sigma^2}{2}\right)\frac{T}{2}} e^{\frac{\sigma}{T} \int_0^T W^{(Q)}(t) dt}$$

$$Q = \left(S_0 e^{\left(r - \frac{\sigma^2}{2}\right)\frac{T}{2}} e^{\frac{\sigma}{T} \int_0^T W^{(Q)}(t) dt} - K \right)_+ \quad \text{WHAT IS THE DENSITY OF THIS R.V.?$$

THEOREM 6.6 : For ALL $f: (0, \infty) \rightarrow \mathbb{R}$

$$X(T) = \underbrace{g(T) W(T)} - \underbrace{\int_0^T g'(s) W(s) ds} \in N(0, \Delta(T))$$

$$\Delta(T) = \int_0^T g(s)^2 ds$$

TO CHOOSE g IN
WE WANT THIS THEOREM TO DERIVE THE
DENSITY OF $\int_0^T W(t) dt$. WE NEED

$$g'(t) = -1 \quad g(T) = 0 \Rightarrow g(t) = T - t$$

HENCE $\int_0^T W(t) dt \in N(0, \Delta(T)) = N(0, T^3/3)$

$$\text{WHERE } \Delta(T) = \int_0^T (T-t)^2 dt$$

$$= T^3/3$$

$$Q = \left(S_0 e^{\frac{1}{2}(2 - \frac{\sigma^2}{2})T + \frac{\sigma}{T} G - K} \right)_+ = \frac{\sigma}{T} \sqrt{\frac{T^3}{3}} \frac{G}{\sqrt{\frac{T^3}{3}}} = \frac{\sigma}{T} \sqrt{\frac{T^3}{3}} H$$

$H \in N(0, 1)$

WHERE $G \in N(0, T^3/3)$

FOR THE STANDARD CALL

$$Y = \left(S_0 e^{(2 - \sigma^2/2)T + \sqrt{T} H} - K \right)_+$$

$\sigma W(T) = \sigma \sqrt{T} \left(\frac{W(T)}{\sqrt{T}} \right) = \sigma \sqrt{T} H$

$H \in N(0, 1)$

$$D = \left(S_0 e^{\left(\frac{1}{2} \left(r - \frac{\sigma^2}{2} \right) T + \sigma \sqrt{\frac{T}{3}} H \right) - K} \right)_+$$

HENCE THE PRICE CAN BE
COMPUTED AS FOR THE STANDARD
CALL