

# Spatial Statistics and Image Analysis

## Lecture 8

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# Lecture's contents

Today's lecture will cover

- ▶ First/second-order properties of point processes
- ▶ Summary functions
- ▶ Poisson point process
- ▶ Log-Gaussian Cox process
- ▶ Neyman-Scott processes
- ▶ Matérn inhibition processes
- ▶ Pairwise interaction point processes

# Definition: Point processes

A point process  $N$  is a stochastic mechanism or rule to produce point patterns or realisations according to the distribution of the process.

A marked point process is a point process where each point  $x_i$  of the process is assigned a quantity  $m(x_i)$ , called a mark. Often, marks are integers or real numbers but more general marks can also be considered.

# Two interpretations

- ▶  $N$  is a counting measure. For a subset  $B$  of  $\mathbb{R}^d$ ,  $N(B)$  is the random number of points in  $B$ . It is assumed that  $N(B) < \infty$  for all bounded sets  $B$ , i.e. that  $N$  is locally finite.
- ▶  $N$  is a random set, i.e. the set of all points  $x_1, x_2, \dots$  in the process. In other words

$$N = \{x_i\} \text{ or } N = \{x_1, x_2, \dots\}$$

Therefore,  $x \in N$  means that the point  $x$  is in the set  $N$ . The set  $N$  can be finite or infinite. If it is finite the total number of points can be deterministic or random.

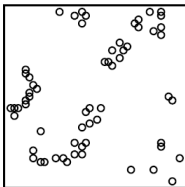
**Remark 1:** We assume that all point processes are simple, i.e. that there are no multiple points ( $x_i \neq x_j$  if  $i \neq j$ ).

**Remark 2:** There is a large literature on processes  $\{Z(t) : t \in T\}$ , where  $T$  is a point process in time. There is an overlap of methods for point processes in space and in time but the temporal case is **not** only a special case of the spatial process with  $d = 1$ . Time is 1-directional.

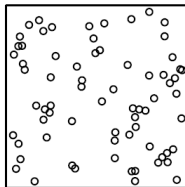
**Remark 3:** To avoid confusion between points of the process and point of  $\mathbb{R}^d$ , the points of the process or point pattern (realization) are called events (or trees or cells).

# Spatial point patterns

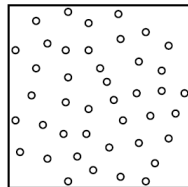
clustered



completely random



regular



- ▶ One of the main questions for point pattern data is usually to determine if we have clustering or repulsion.
- ▶ The completely random case corresponds to the Poisson point process

# Examples

- ▶ Locations of betacells within a rectangular region in a cat's eye (regular)
- ▶ Locations of Finnish pine saplings (clustered)
- ▶ Locations of Spanish towns (regular)
- ▶ Locations of galaxies (clustered)

**Remark:** Very different scales, from microscopic to cosmic

# First-order properties (without marks)

The mean number of points of  $N$  in  $B$  is  $\mathbb{E}(N(B))$  (depends on the set  $B$ ). We use the notation

$$\Lambda(B) = \mathbb{E}(N(B))$$

and call  $\Lambda$  the intensity measure.

Under some continuity conditions, a density function  $\lambda$ , called the intensity function, exists, and

$$\Lambda(B) = \int_B \lambda(x) dx.$$



# Some properties of point processes: stationarity and isotropy

A point process  $N$  is stationary (translation invariant) if  $N$  and the translated point process  $N_x$  have the same distribution for all translations  $x$ , i.e.

$$N = \{x_1, x_2, \dots\} \text{ and } N_x = \{x_1 + x, x_2 + x, \dots\}$$

have the same distribution for all  $x \in \mathbb{R}^d$ .

A point process is isotropic (rotation invariant) if its characteristics are invariant under rotations, i.e.

$$N = \{x_1, x_2, \dots\} \text{ and } rN_x = \{rx_1, rx_2, \dots\}$$

have the same distribution for any rotation  $r$  around the origin.  
If a point process is both stationary and isotropic, it is called motion-invariant.

# First-order properties

If  $N$  is stationary, then

$$\Lambda(B) = \lambda|B|,$$

where  $0 < \lambda < \infty$  is called the intensity of  $N$  and  $|B|$  is the volume of  $B$ .

$\lambda$  is the mean number of points of  $N$  per unit area, i.e.

$$\lambda = \frac{\Lambda(B)}{|B|} = \frac{\mathbb{E}(N(B))}{|B|}.$$

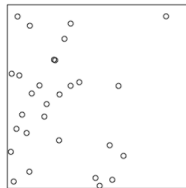
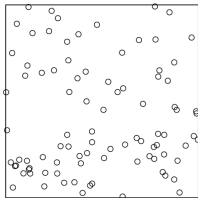
# Homogeneous Poisson point process

A point process is a homogeneous Poisson process (CSR) if

- (P1) for some  $\lambda > 0$  and any finite region  $B$ ,  $N(B)$  has a Poisson distribution with mean  $\lambda|B|$
- (P2) given  $N(B) = n$ , the events in  $B$  form an independent random sample from the uniform distribution on  $B$

Inhomogeneous Poisson process: intensity  $\lambda$  (in homogeneous Poisson process) replaced by an intensity function  $\lambda(x)$

# Examples: realizations of Poisson processes



**Left:** Homogeneous Poisson point process with intensity 100

**Right:** A Poisson point process with a spatially varying intensity function  $\lambda(x, y) = 100e^{-3x}$

## Some comments

- ▶ Despite its simplicity the Poisson point process has a fundamental role in the characterisation of the spatial structure of spatial point processes.
- ▶ It is used as a **reference model** to assess clustering or inhibition.
- ▶ It is often used as a **building block** for constructing more complex point processes, for example, Neyman-Scott point processes (clustered) and Matérn hard-core point processes (regular).

# Two distribution functions

1. Let  $D_1$  denote the distance from an arbitrary event  $x$  to the nearest other event. Then, the nearest neighbour distance function is

$$G(r) = P(D_1 \leq r) = 1 - P(N(b(x, r) \setminus \{x\}) = 0)$$

If the pattern is completely spatially random (CSR),  
 $G(r) = 1 - \exp(-\lambda\pi r^2)$ . For regular patterns  $G(r)$  tends to lie below and for clustered patterns above the CSR curve.

2. Let  $D_2$  denote the distance from an arbitrary point  $x$  to the nearest event. Then,

$$F(r) = P(D_2 \leq r) = 1 - P(N(b(x, r)) = 0)$$

If the pattern is completely spatially random,  
 $F(r) = 1 - \exp(-\lambda\pi r^2)$ . For regular patterns  $F(r)$  tends to lie above and for clustered patterns below the CSR curve.

# Combination of the two

Using  $G$  and  $F$  we can define the so-called  $J$  function as

$$J(r) = \frac{1 - G(r)}{1 - F(r)}$$

(whenever  $F(r) > 0$ )

If the pattern is completely spatially random,  $J(r) \equiv 1$ . For regular patterns  $J(r) > 1$  and for clustered patterns  $J(r) < 1$ .

## Second-order properties

The 2nd order properties of a stationary and isotropic point process can be characterized by Ripley's  $K$  function

$$K(r) = \lambda^{-1} \mathbb{E}[\# \text{ further events within distance } r \text{ of a typical event}].$$

Often, a transformed and centered version of the  $K$  function is used,

$$L(r) - r = \sqrt{K(r)/\pi} - r,$$

which equals 0 under CSR. Values less than zero indicate regularity and values larger than zero clustering.



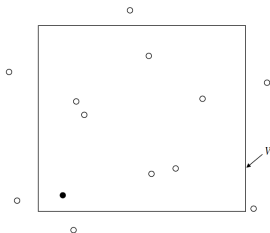
# Estimation of the K function.

- An naive estimate for the  $K$  function is given by

$$\hat{K}(r) = \frac{1}{n\hat{\lambda}} \sum_{i=1}^n \sum_{i \neq j} 1\{\|x_i - x_j\| \leq r\}$$

where  $\hat{\lambda} = \frac{n-1}{|W|}$  is an estimate for  $\lambda$

- Typically, a point pattern is observed in a (bounded) observation window and points outside the window are not observed. Hence this estimator is biased.



# Edge corrections

- ▶ Estimators of the summary functions need to be edge-corrected
- ▶ An unbiased estimate for the  $K$  function is given by

$$\hat{K}(r) = \frac{1}{n\hat{\lambda}} \sum_{i=1}^n \sum_{i \neq j} w(x_i, x_j) 1\{\|x_i - x_j\| \leq r\}$$

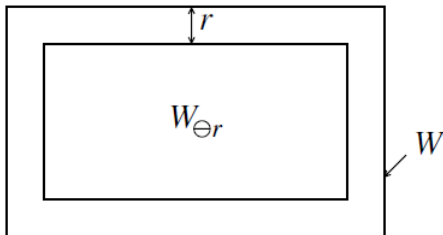
where  $w(x_i, x_j)$  is an edge correction term.

- ▶ Edge correction methods include minus sampling or border method, Ripley's isotropic correction and translation (stationary) correction

## Minus sampling/ border correction.

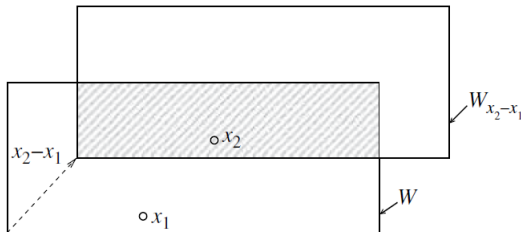
- Let  $W_{\ominus r}$  denote the subset of  $W$  the points in which are in the interior of  $W$  and have a distance larger than  $r$  from the boundary  $\partial W$ . We consider pair of points  $x \in W_{\ominus r}$  and  $y \in W$ . Then an unbiased estimate for  $K$  is given by

$$\hat{K}_{minus}(r) = \frac{1}{\hat{\lambda}^2 |W_{\ominus r}|} \sum_{x \in X \cap W_{\ominus r}} \sum_{y \in X} 1\{\|x - y\| \leq r\}$$



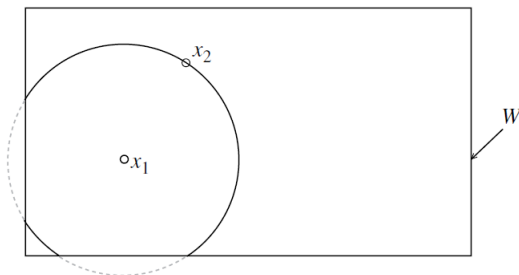
# Translational edge-correction

- ▶ Applicable for stationary point processes
- ▶ The weights  $w(x_i, x_j) = \frac{1}{|W \cap W_{x_i - x_j}|}$  are given by the area of the intersection of  $W$  with the translated by  $x_i - x_j$  window  $W_{x_i - x_j}$ .

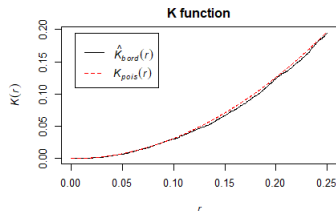
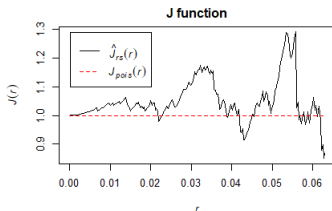
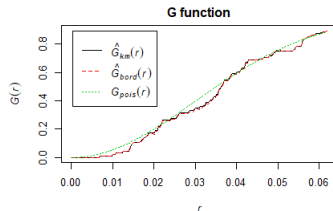
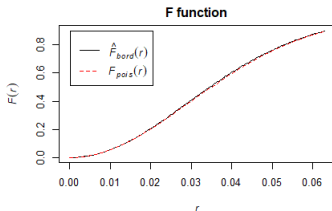


# Isotropic edge-correction

- ▶ Applicable for stationary and isotropic point processes
- ▶ The weights  $w(x_i, x_j) = \frac{\nu_1(\partial b(x_i, \|x_i - x_j\|) \cap W)}{2\pi \|x_i - x_j\|}$  where  $\nu_1$  denotes the length of a curve,  $\partial$  denote the boundary of a set and  $b(x_i, r)$  the ball centred at  $x_i$  with radius  $r$ .
- ▶ The weights give the proportion of the circle that lies in  $W$ .

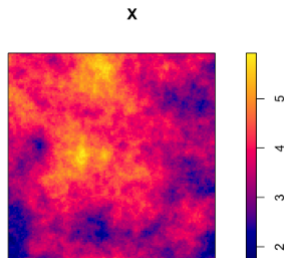
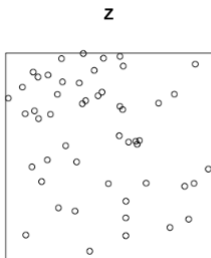


# Summary statistics for CSR



# Log-Gaussian Cox process

- Hierarchical model, where  $X$  is a Gaussian random field and  $Z | X$  is an inhomogeneous Poisson process where  $\lambda(x, y) = \exp(X(x, y))$
- Example:  $X$  is a Gaussian random field with mean 3 and an exponential covariance function.



# Neyman-Scott processes

Cluster processes are models for aggregated spatial point patterns

For Neyman-Scott cluster process

- (MC1) parent events form a Poisson process with intensity  $\lambda$
- (MC2) each parent produces a random number  $S$  of daughters (offsprings), realized independently and identically for each parent according to some probability distribution  $p_s$
- (MC3) the locations of the daughters in a cluster are independently and identically distributed according to a bivariate continuous probability density function.

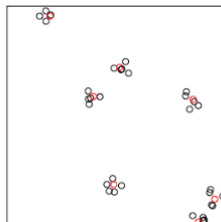
The cluster process consists only of the daughter points.



# Examples of Neyman-Scott processes

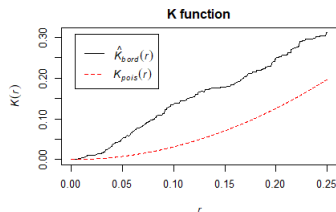
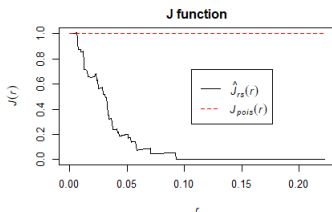
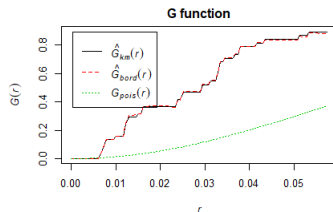
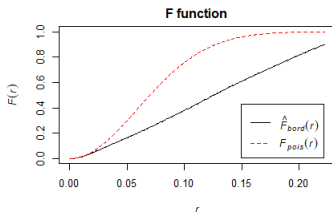
- ▶ Matérn cluster process:  $p_s$  is a Poisson distribution and the continuous distribution for the locations of offspring is the uniform distribution on a disc
- ▶ Thomas cluster process:  $p_s$  is a Poisson distribution and the continuous distribution for the locations of offspring is the the 2-dimensional normal distribution

# Realization of a Matérn cluster process



**Red:** Parent points from a Poisson point process with intensity 7  
**Black:** Daughter points with cluster radius 0.05 and average number of daughter per cluster 5.

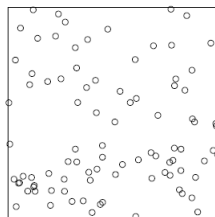
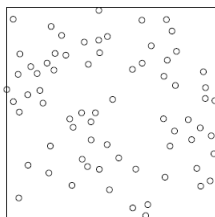
# Summary statistics



# Matérn I hard-core process

- ▶ Hard-core processes are models for regular spatial point patterns
- ▶ Matérn I hard-core process:
  1. Simulate a homogeneous Poisson process  $\mathcal{Z}$
  2. Delete any point in  $\mathcal{Z}$  that lies closer than a distance  $r$  from the nearest other point
- ▶ There is a minimum allowed distance, called hard-core distance, between any two points
- ▶ Matérn I hard-core process: A Poisson process with intensity  $\lambda$  is thinned by deleting all pairs of points that are at distance less than the hard-core radius apart.

# Realization of a Matérn I hard-core process

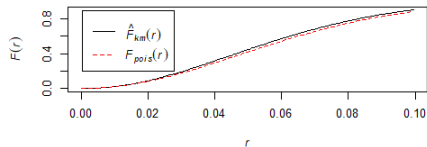


**Left:** Hard-core process with the initial Poisson intensity 300, hard-core radius 0.04

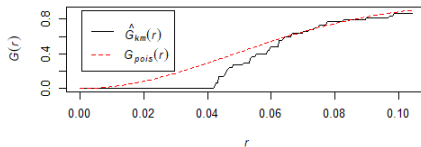
**Right:** Poisson process with intensity 100

# Summary statistics

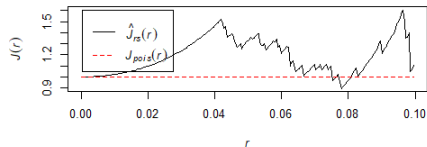
F function



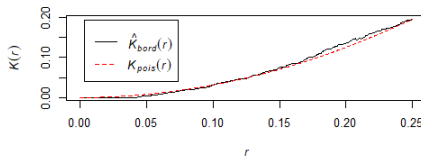
G function



J function



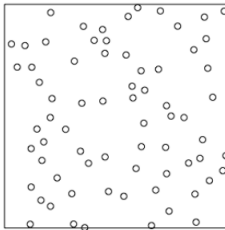
K function



# Matérn II hard-core process

The thinning strategy for Matérn II hard-core processes is the following

1. Simulate a homogeneous Poisson process  $Z$
2. Mark each point in  $Z$  by “ages”, which are independent and uniformly distributed numbers in  $[0,1]$ .
3. Delete any point in  $Z$  that lies closer than a distance  $r$  from another point that has a higher age.



# Pairwise interaction processes

- ▶ Pairwise interaction processes are a subclass of Markov point processes which are models for point patterns with interaction between the events
- ▶ There is interaction between the events if they are "neighbours", e.g. if they are close enough to each other
- ▶ Models for inhibition/regularity



# Pairwise interaction processes: Strauss process

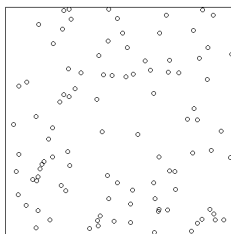
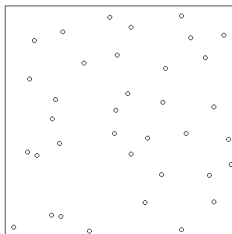
- ▶ Two points are neighbours if they are closer than distance  $R$  apart
- ▶ The density function (with respect to a Poisson process with intensity 1) is

$$f(x) = \alpha \beta^{n(x)} \gamma^{s(x)}, \quad \beta > 0, \quad \gamma \geq 0,$$

where

- ▶  $\beta > 0$  is the effect of a single event (connected to the intensity of the process)
- ▶  $0 < \gamma \leq 1$  is an interaction parameter
- ▶  $n(x)$  is the number of points in the configuration
- ▶  $s(x)$  is the number of  $R$  close pairs in the configuration, where  $R > 0$  is an interaction radius (range of interaction)
- ▶  $\alpha$  is a normalizing constant

## Example: Strauss process



Left: Strauss process with  $\beta = 100$ ,  $\gamma = 0.2$  and  $R = 0.1$

Right: Strauss process with  $\beta = 100$ ,  $\gamma = 1$  and  $R = 0.1$

- ▶  $\gamma = 1$  corresponds to CSR
- ▶  $\gamma < 1$  corresponds to inhibition

# Simulation based envelope tests

- ▶ Let  $\hat{K}(r)$  be an estimate of the  $K$  function for  $X$ .
- ▶ Comparing  $\hat{K}(r)$  with the theoretical value under CSR ( $K(r) = \pi r^2$ ) we can get information about whether there is clustering or inhibition at scale  $r$ .
- ▶ But how large deviation could we expect to find by pure randomness?
- ▶ Simulation based envelope techniques have been developed to tackle this issue.

# Simulation based envelope tests

- ▶ Let  $N(W) = n$ , then we
  1. Simulate  $X_1, \dots, X_M$  on  $W$  from a Poisson process conditioned on  $N_{X_i}(W) = n$  (Binomial point process).
  2. Calculate  $\hat{K}_i(r)$  for all  $X_i$  and check if

$$\min_i \hat{K}_i(r) \leq \hat{K}(r) \leq \max_i \hat{K}_i(r)$$

- ▶ For a **fixed value of  $r$**  and for  $M = 39$  this is a valid test at level  $\alpha = 0.05$ .
- ▶ However, the probability that the  $K$  function will wander outside the envelope for some  $r$  is very much higher than 0.05.
- ▶ Recently, statistically valid envelope tests have been developed ([Myllymäki et al. 2017](#))

- ▶ Myllymäki, M., Mrkvička, T., Grabarnik, P., Seijo, H., Hahn, U. (2017). Global envelope tests for spatial processes. *Journal of the Royal Statistical Society: Series B (Statistical Methodology)*, 79(2), 381-404.
- ▶ Baddeley, A., Turner, R. Spatstat: an R package for analyzing spatial point patterns. *J. Stat. Softw.* 12 (2005) 1-42.
- ▶ Besag, J.E., 1977. Comment on “Modelling spatial patterns” by B. D. Ripley. *Journal of the Royal Statistical Society B (Methodological)* 39 (1977) 193-195.
- ▶ Illian, J., Penttinen, A., Stoyan, H., Stoyan, D. *Statistical Analysis and Modelling of Spatial Point Patterns*. Chichester: Wiley (2008).
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