

MVE550 2022 Lecture 10

Poisson processes

Dobrow chapter 6, sections 6.6 - 6.7

Continuous-time Markov chains

Dobrow chapter 7, sections 7.1 - 7.4

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Poisson processes: Short review from last lecture

- ▶ Three different but equivalent definitions.
- ▶ Some selected properties:
 - ▶ $N_t \sim \text{Poisson}(\lambda t)$
 - ▶ Independent increments
 - ▶ Stationary increments
 - ▶ Inter-arrival times distributed as $X_i \sim \text{Exponential}(\lambda)$.
 - ▶ Arrival times distributed as $S_n \sim \text{Gamma}(n, \lambda)$.
 - ▶ Superposition and thinning.
- ▶ Another important property: Let $(N_t)_{t \geq 0}$ be a Poisson process with parameter λ . If we fix that $N_t = k$ and we select one of these k arrivals, its arrival time is uniformly distributed on the interval $[0, t]$.

Spatial Poisson processes

- ▶ A collection of random variables $\{N_A\}_{A \subseteq \mathbb{R}^d}$ is a spatial Poisson process with parameter λ if
 - ▶ For each bounded set $A \subseteq \mathbb{R}^d$, N_A has a Poisson distribution with parameter $\lambda|A|$.
 - ▶ Whenever $A \subseteq B$, $N_A \leq N_B$.
 - ▶ Whenever A and B are disjoint sets, N_A and N_B are independent.
- ▶ Simulate by first simulating the total (Poisson distributed) and then place points independently uniformly within the area.
- ▶ One may use simulations to estimate properties such as the average distance to the nearest neighbour (or the third nearest neighbour or whatever).
- ▶ Quite useful model in practice.

Non-homogeneous Poisson processes

- ▶ A counting process $\{N_t\}_{t \geq 0}$ is a *non-homogeneous* Poisson process with intensity function $\lambda(t)$ if
 - ▶ $N_0 = 0$.
 - ▶ For $t > 0$,

$$N_t \sim \text{Poisson} \left(\int_0^t \lambda(x) dx \right)$$

- ▶ It has independent increments.
- ▶ Again a very flexible and useful model in practice.
- ▶ One may have non-homogeneous spatial Poisson processes.

Introduction to continuous-time Markov chains

- ▶ We now consider general continuous-time discrete state space Markov chains.
- ▶ Comparing to counting processes: We can now potentially jump between any two states.
- ▶ Comparing to the discrete-time Markov chains: We now model that we stay in each state for some real-valued amount of time.
- ▶ The Markov property is a type of “memorylessness”: The property will imply that the amount of time spent in each state is Exponentially distributed.
- ▶ Very useful tool, can be used to model for example queues.

Example

- ▶ We have previously discussed modelling the weather as a discrete time Markov chain where the weather *each day* is “rain”, “snow”, or “clear”, with transition matrix for example

$$P = \begin{bmatrix} 0.2 & 0.6 & 0.2 \\ 0.1 & 0.8 & 0.1 \\ 0.1 & 0.4 & 0.5 \end{bmatrix}.$$

- ▶ A more realistic model is that each weather type lasts some length of time, before changing to a *different* weather type:
 - ▶ Let's say the time each weather type lasts is Exponentially distributed with parameters q_r , q_s and q_c (so that expected durations of weather types are $1/q_r$, $1/q_s$, $1/q_c$, respectively).
 - ▶ Transitions after this time could happen according to a transition matrix, for example

$$\tilde{P} = \begin{bmatrix} 0 & 3/4 & 1/4 \\ 1/2 & 0 & 1/2 \\ 1/5 & 4/5 & 0 \end{bmatrix}.$$

- ▶ Note that the process is completely described by parameters q_r, q_s, q_c and p_{ij} , where $\tilde{P}_{ij} = p_{ij}$. Note that $p_{ii} = 0$ for all i .

Continuous time Markov chains

- ▶ A continuous time stochastic process $\{X_t\}_{t \geq 0}$ with discrete state space S is a *continuous time Markov chain* if

$$P(X_{t+s} = j \mid X_s = i, X_u, 0 \leq u < s) = P(X_{t+s} = j \mid X_s = i)$$

where $s, t \geq 0$ and $i, j, x_u \in S$.

- ▶ The process is *time-homogeneous* if for $s, t \geq 0$ and all $i, j \in S$

$$P(X_{t+s} = j \mid X_s = i) = P(X_t = j \mid X_0 = i)$$

.

- ▶ We then define the *transition function* as the matrix function $P(t)$ with the entries of the matrix given by

$$P(t)_{ij} = P(X_t = j \mid X_0 = i)$$

The Chapman-Kolmogorov Equations

For the transition function $P(t)$ we have

- ▶ $P(s + t) = P(s)P(t)$ (Note: Matrix equation!)
- ▶ $P(0) = I$
- ▶ Note similarity to the properties of the exponential function!
However, $P(t)$ is a matrix, not a number.
- ▶ Example:
 - ▶ A Poisson process with parameter λ is a continuous time time-homogeneous Markov chain.
 - ▶ We get

$$P(t) = \begin{bmatrix} e^{-\lambda t} & (\lambda t)e^{-\lambda t} & (\lambda t)^2 e^{-\lambda t}/2! & (\lambda t)^3 e^{-\lambda t}/3! & \dots \\ 0 & e^{-\lambda t} & (\lambda t)e^{-\lambda t} & (\lambda t)^2 e^{-\lambda t}/2! & \dots \\ 0 & 0 & e^{-\lambda t} & (\lambda t)e^{-\lambda t} & \dots \\ 0 & 0 & 0 & e^{-\lambda t} & \dots \\ \vdots & \vdots & \vdots & \vdots & \ddots \end{bmatrix}$$

Holding times are exponentially distributed

- ▶ Define T_i as the time the continuous-time Markov chain started in i stays in i before moving to a different state, so that for any $s > 0$

$$P(T_i > s) = P(X_u = i, 0 \leq u \leq s)$$

- ▶ The distribution of T_i is *memoryless* and thus exponential.
- ▶ We define q_i so that

$$T_i \sim \text{Exponential}(q_i)$$

- ▶ Remember that this means that the average time the process stays in i is $1/q_i$. The *rate* of transition out of the state is q_i .
- ▶ Note that we can have $q_i = 0$ meaning that the state i is *absorbing*: $P(T_i > s) = 1$.

The embedded chain

- ▶ Define a new stochastic process by listing the states the chain visits. This will be a discrete time Markov chain.
- ▶ It is called the *embedded chain*; transition matrix is denoted $\tilde{P}$.
- ▶ Note that $\tilde{P}$ has zeros along its diagonal!
- ▶ Note that the continuous time Markov chain is completely determined by the expected holding times $(1/q_1, \dots, 1/q_k)$ and the transition matrix $\tilde{P}$ of the embedded chain. We write p_{ij} for the entries of $\tilde{P}$.

Describing the chain using transition rates

A way to describe a continuous-time Markov chain is to describe $k \times (k - 1)$ independent “alarm clocks”:

- ▶ For states i and j so that $i \neq j$, let q_{ij} be the parameter of an Exponentially distributed random variable representing the time until an “alarm clock” rings.
- ▶ When in state i , wait until the *first* alarm clock rings, then move to the state given by the index j of that alarm clock. This defines a continuous-time Markov chain.
- ▶ The time until the first alarmclock rings is Exponentially distributed with parameter given by

$$q_i = q_{i1} + q_{i2} + \cdots + q_{i,i-1} + q_{i,i+1} + \cdots + q_{ik} \quad (1)$$

i.e., the parameter of the holding time distribution at i .

- ▶ The chain is completely described by the rates q_{ij} , $i \neq j$.
- ▶ We saw above: The chain is also completely determined by the p_{ij} and the q_i . The relationship is described by Equation 1 and, for $i \neq j$,

$$p_{ij} = \frac{q_{ij}}{q_{i1} + q_{i2} + \cdots + q_{i,i-1} + q_{i,i+1} + \cdots + q_{ik}} = \frac{q_{ij}}{q_i}.$$

The derivative of $P(t)$ at zero

- ▶ To relate $P(t)$ to the q_{ij} 's, we first relate them to $P'(0)$.
- ▶ Assuming $P(t)$ is differentiable we can show that

$$P'(0) = \begin{bmatrix} -q_1 & q_{12} & q_{13} & \dots & q_{1k} \\ q_{21} & -q_2 & q_{23} & \dots & q_{2k} \\ q_{31} & q_{31} & -q_3 & \dots & q_{3k} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ q_{k1} & q_{k2} & q_{k3} & \dots & -q_k \end{bmatrix} = Q$$

where the q_i and the q_{ij} are those defined earlier.

- ▶ Note that the rows of $P'(0)$, i.e., Q , sum to zero!
- ▶ In fact we don't need to require a finite state space; discrete is enough.
- ▶ Q is called the (*infinitesimal*) *generator* of the chain.

Kolmogorov Forward Backward

- ▶ Prove: We get that for all $t \geq 0$,

$$P'(t) = P(t)Q = QP(t)$$

- ▶ Note what this means in terms of the components of $P(t)$:

$$P'(t)_{ij} = -P_{ij}(t)q_j + \sum_{k \neq j} P_{ik}(t)q_{kj}$$

$$P'(t)_{ij} = -q_i P_{ij}(t) + \sum_{k \neq i} q_{ik} P_{kj}(t)$$

- ▶ Either line with equations above define a set of differential equations which the components of the matrix function $P(t)$ needs to fulfill.

The matrix exponential

- ▶ For any square matrix A define the *matrix exponential* as

$$e^A = \sum_{n=0}^{\infty} \frac{1}{n!} A^n = I + A + \frac{1}{2}A^2 + \frac{1}{6}A^3 + \frac{1}{24}A^4 + \dots$$

- ▶ The series converges for all square matrices A (we don't show this).
- ▶ Some important properties:
 - ▶ $e^0 = I$.
 - ▶ $e^A e^{-A} = I$.
 - ▶ $e^{(s+t)A} = e^{sA} e^{tA}$.
 - ▶ If $AB = BA$ then $e^{A+B} = e^A e^B = e^B e^A$.
 - ▶ $\frac{\partial}{\partial t} e^{tA} = A e^{tA} = e^{tA} A$.
- ▶ $P(t) = e^{tQ}$ is the unique solution to the differential equations $P'(t) = QP(t)$ for all $t \geq 0$ and $P(0) = I$.
- ▶ In R you may use `expm` from R package `expm` to compute exponential matrices.

Computing the matrix exponential

- Assume there exists an invertible matrix S and a matrix D such that $Q = SDS^{-1}$. Then (show!)

$$e^{tQ} = Se^{tD}S^{-1}$$

- If $D = \begin{bmatrix} \lambda_1 & 0 & \dots & 0 \\ 0 & \lambda_2 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \lambda_k \end{bmatrix}$ is a diagonal matrix, then (show!)

$$e^{tD} = \begin{bmatrix} e^{t\lambda_1} & 0 & \dots & 0 \\ 0 & e^{t\lambda_2} & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & e^{t\lambda_k} \end{bmatrix}.$$

- Recall that if Q is *diagonalizable* it can be written as $Q = SDS^{-1}$ where D is diagonal with the eigenvalues along the diagonal, and S has the corresponding eigenvectors as columns.

Limiting and stationary distributions

- ▶ A probability vector v represents a *limiting distribution* if, for all states i and j ,

$$\lim_{t \rightarrow \infty} P_{ij}(t) = v_j.$$

- ▶ A probability vector v represents a *stationary distribution*, if, for all $t \geq 0$,

$$v = vP(t)$$

- ▶ Note: This happens if and only if $0 = vQ$.
- ▶ A limiting distribution is a stationary distribution but not necessarily vice versa.
- ▶ A continuous-time Markov chain is *irreducible* if for all i and j there exists a $t > 0$ such that $P_{ij}(t) > 0$.
- ▶ However, periodic continuous-time Markov chains do not exist: If $P_{ij}(t) > 0$ for some $t > 0$ then $P_{ij}(t) > 0$ for all $t > 0$.

The fundamental limit theorem

- ▶ An absorbing communication class is one where there is zero probability (i.e., zero rate) of leaving it to other communication classes.
- ▶ For a finite-state continuous-time Markov chain (with finite holding time parameters) there are two possibilities:
 - ▶ The process is irreducible, and $P_{ij}(t) > 0$ for all $t > 0$ and all i, j .
 - ▶ The process contains one or more absorbing communication classes.
- ▶ **Fundamental Limit Theorem:** Let $\{X_t\}_{t \geq 0}$ be a finite, irreducible, continuous-time Markov chain with transition function $P(t)$. Then there exists a unique stationary distribution vector v which is also the limiting distribution.
- ▶ The limiting distribution of such a chain can be found as the unique v satisfying $vQ = 0$.