

# MVE550 2023 Lecture 15

## Dobrow Chapter 7, part 3

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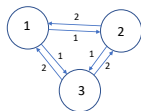
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December 7, 2023

## Review from last time:

- ▶ Continuous-time Markov chains, described by a generator matrix  $Q$ , consisting of rates.
- ▶  $P(t) = e^{tQ}$ .
- ▶ Limiting and stationary distributions, the fundamental limit theorem.
- ▶ Solve  $vQ = 0$  to find limiting distribution.
- ▶ Absorbing states, the fundamental matrix  $F = -V^{-1}$ .
- ▶ The stationary distribution of a continuous-time chain and its embedded chain are different, but related.

# Global Balance



- ▶ Let  $v = (v_1, v_2, v_3)$  be the stationary distribution. At  $v$ , the flow into a state must be equal to the flow out of that state.
- ▶ We get:  $2v_1 = 2v_2 + 2v_3$ ,  $3v_2 = 1v_1 + 2v_3$ , and  $4v_3 = 1v_1 + 1v_2$ .
- ▶ Note that these are exactly the equations we get from  $vQ = 0$ :

$$(v_1, v_2, v_3) \begin{bmatrix} -2 & 1 & 1 \\ 2 & -3 & 1 \\ 2 & 2 & -4 \end{bmatrix} = 0$$

- ▶ This happens because  $vQ = 0$  gives for each state  $j$

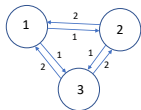
$$\sum_{i \neq j} v_i q_{ij} = v_j q_{jj}$$

- ▶ These are called the global balance equations.
- ▶ Generalization (proof in Dobrow): If  $A$  is a set of states, then the long term rates of movement *into* and *out of*  $A$  are the same:

$$\sum_{i \notin A} \sum_{j \in A} v_i q_{ij} = \sum_{i \notin A} \sum_{j \in A} v_j q_{ji}$$

# Local balance and time reversibility

- ▶ Local balance: A stronger condition: The flow between *every pair* of states is balanced. This is *not* true for all models!
- ▶ For the model above the local balance equations are



$$1v_1 = 2v_2, 1v_2 = 2v_3, \text{ and } 2v_3 = 1v_1.$$

- ▶ These are *not* satisfied by the stationary  $v$ ! (Check it).
- ▶ An irreducible continuous-time Markov chain with stationary distribution  $v$  is said to be *time reversible* if for all  $i, j$ ,

$$v_i q_{ij} = v_j q_{ji}$$

which is in fact the *local balance* condition.

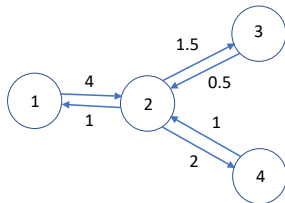
- ▶ Note: The rate of observed changes from  $i$  to  $j$  is the same as the rate of observed changes from  $j$  to  $i$ . Thus this is called time reversibility.
- ▶ Note that (similar to discrete chains): If a probability vector  $v$  satisfies local balance condition, then  $v$  is a stationary distribution. (Easy to show directly).

# Markov processes with transition graphs that are trees

- ▶ **A tree is a graph that does not contain cycles.**
- ▶ Assume the transition graph of an irreducible continuous-time Markov chain is a tree.
- ▶ In a tree, any edge between two states divides all states into two groups (each on each side of the edge). Thus, the flow must be balanced across each edge.
- ▶ It follows that the Markov chain must satisfy the local balance condition, i.e., be time reversible, i.e.,  $v_i q_{ij} = v_j q_{ji}$  for all  $i$  and  $j$ .
- ▶ More formally, this can be proved using the generalized global balance property.
- ▶ Note that the process can be time reversible even if the transition graph is not a tree.

# Example

- Consider the continuous-time Markov chain with transition graph



- As the transition graph is a tree, the chain is necessarily time reversible. We can find the stationary distribution by considering the local balance equations:

$$4v_1 = 1v_2, \quad 1.5v_2 = 0.5v_3, \quad 2v_2 = 1v_4$$

- Together with the equation  $v_1 + v_2 + v_3 + v_4 = 1$  we easily get the limiting distribution

$$v = \left( \frac{1}{25}, \frac{4}{25}, \frac{12}{25}, \frac{8}{25} \right)$$

# Birth-and-death processes

- ▶ A birth-and-death process is a continuous-time Markov chain where the state space is the set of nonnegative integers and transitions only occur to neighbouring integers.
- ▶ The process is necessarily time-reversible, as the transition graph is a tree (in fact, a line).
- ▶ We denote the rate of *births* from  $i$  to  $i + 1$  with  $\lambda_i$ , and the rate of *deaths* from  $i$  to  $i - 1$  with  $\mu_i$ .
- ▶ The generator matrix is

$$Q = \begin{bmatrix} -\lambda_0 & \lambda_0 & 0 & 0 & \dots \\ \mu_1 & -(\mu_1 + \lambda_1) & \lambda_1 & 0 & \dots \\ 0 & \mu_2 & -(\mu_2 + \lambda_2) & \lambda_2 & \dots \\ 0 & 0 & \mu_3 & -(\mu_3 + \lambda_3) & \dots \\ \vdots & \vdots & \vdots & \vdots & \ddots \end{bmatrix}$$

- ▶ Provided  $\sum_{k=1}^{\infty} \prod_{i=1}^k \frac{\lambda_{i-1}}{\mu_i} < \infty$ , the unique stationary distribution is given by

$$v_k = v_0 \prod_{i=1}^k \frac{\lambda_{i-1}}{\mu_i} \text{ for } k = 1, 2, \dots, \quad v_0 = \left( 1 + \sum_{k=1}^{\infty} \prod_{i=1}^k \frac{\lambda_{i-1}}{\mu_i} \right)^{-1}$$

## Example

- ▶ The simplest example of a birth-and-death process is one where all birth rates  $\lambda_i$  and all death rates  $\mu_i$  are the same values  $\lambda$  and  $\mu$ , respectively.
- ▶ We get that

$$v_k = v_0 \prod_{i=1}^k \frac{\lambda}{\mu} = v_0 \left( \frac{\lambda}{\mu} \right)^k$$

$$v_0 = \left( \sum_{k=0}^{\infty} \left( \frac{\lambda}{\mu} \right)^k \right)^{-1} = \frac{1}{1 + \frac{\lambda}{\mu} + \left( \frac{\lambda}{\mu} \right)^2 + \dots} = \frac{1}{1/(1 - \frac{\lambda}{\mu})} = 1 - \frac{\lambda}{\mu}$$

- ▶ We see that the limiting distribution is Geometric  $\left(1 - \frac{\lambda}{\mu}\right)$ .
- ▶ For example, the long-term average value of  $X_t$  will be

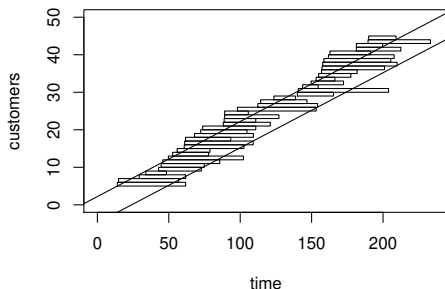
$$\frac{\lambda/\mu}{1 - \lambda/\mu} = \frac{\lambda}{\mu - \lambda}$$



# Queueing theory

- ▶ Birth-and-death processes are special cases of queues.
- ▶ In the more general theory of queues:
  - ▶ The arrival process (“births”) need not be a Poisson process, with exponentially distributed inter-arrival times.
  - ▶ The service times in the system need not be exponentially distributed.
  - ▶ There can be many other generalizations, such as how many servers there are, how the servers work, how the lines work, etc.
- ▶ One can use notation  $A/B/n$  where  $A$  denotes arrival process,  $B$  denotes service process, and  $n$  the number of servers.
- ▶ With this notation, our birth-and-death model above with constant birth and death rates is denoted  $M/M/1$ . (M means Markov).
- ▶ Generalize our results to the case  $M/M/c$  by taking into account that some servers are idle when there are fewer than  $c$  customers.

# Little's formula



Little's formula is valid when customers arrive at rate  $\lambda$ , and stay an expected time  $W$ . The left line represents the average arrival times of customers: It has slope  $\lambda$ . The right line represents the average departure time of customers. The horizontal distance between the lines is  $W$ . The vertical distance between the lines will be  $L$ , the average number of customers in the system. Thus

$$\lambda = \frac{L}{W}$$

# Poisson subordination

- ▶ We may simulate from a continuous time finite state Markov chain by drawing the holding times from distributions  $\text{Exponential}(q_i)$ , where  $q_i$  depends on the state  $i$ , and then use  $\tilde{P}$ .
- ▶ INSTEAD, always simulate holding times from  $\text{Exponential}(\lambda)$  where  $\lambda$  is large, and allow movement back to the same state.
- ▶ Mathematical formulation: Given generator matrix  $Q$ . If  $\lambda \geq \max(q_1, \dots, q_n)$  then
  - ▶  $R = \frac{1}{\lambda}Q + I$  is a stochastic matrix.
  - ▶ We can write

$$P(t) = e^{tQ} = e^{-t\lambda I} e^{t\lambda R} = e^{-t\lambda} \sum_{k=0}^{\infty} \frac{(t\lambda R)^k}{k!} = \sum_{k=0}^{\infty} R^k \frac{e^{-\lambda t} (\lambda t)^k}{k!}.$$

- ▶ Thus: To find the probability of going from  $i$  to  $j$  during time  $t$ :
  1. Condition on number of changes occurring  $k \sim \text{Poisson}(\lambda t)$ .
  2. For each  $k$ , use  $k$  steps of discrete chain with transition matrix  $R$ .
- ▶ This provides a good way to compute  $e^{tQ}$ : Throw away terms where  $k$  is over some limit. Better accuracy than using definition of exponential matrix!
- ▶ This discrete chain has the same stationary distribution as the continuous chain.