

Probabilistic Approach to Linear Regression

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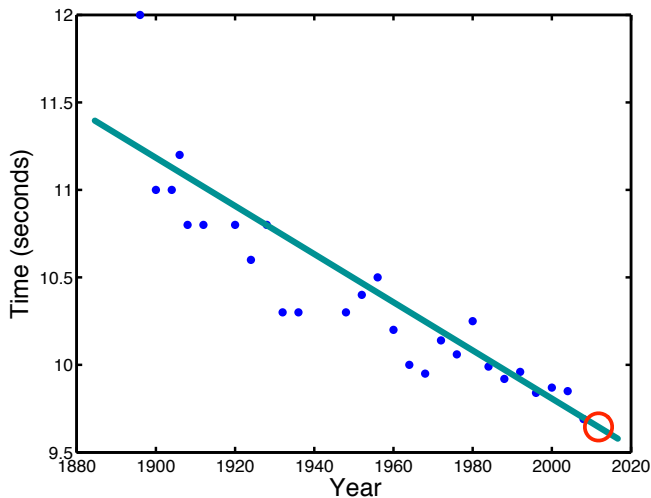
Reference

The content and the slides are adapted from

S. Rogers and M. Girolami, A First Course in Machine Learning (FCML), 2nd edition, Chapman & Hall/CRC 2016, ISBN: 9781498738484

Some data and a problem

Use the model (line) to *predict* the winning time in 2012.



Recipe for a linear model

More complex model: $t = w_0 + w_1x + w_2x^2 + \dots + w_Dx^D$

$$\mathbf{x}_n = \begin{bmatrix} 1 \\ x_n \\ x_n^2 \\ \vdots \\ x_n^D \end{bmatrix}, \mathbf{X} = \begin{bmatrix} 1 & x_1^1 & x_1^2 & \dots & x_1^D \\ 1 & x_2^1 & x_2^2 & \dots & x_2^D \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & x_N^1 & x_N^2 & \dots & x_N^D \end{bmatrix} \mathbf{t} = \begin{bmatrix} t_1 \\ t_n \\ \vdots \\ t_N \end{bmatrix},$$

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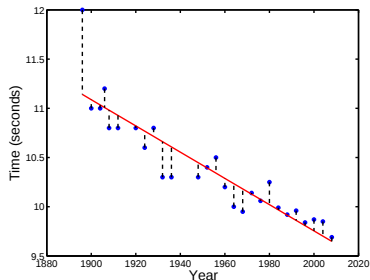
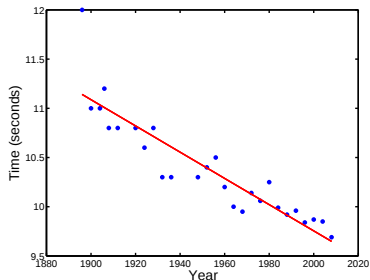
$$\mathbf{w} = \begin{bmatrix} w_0 \\ w_1 \\ \vdots \\ w_D \end{bmatrix}, \quad \text{Model : } t_n = \mathbf{w}^T \mathbf{x}_n, \quad \text{or} \quad \mathbf{t} = \mathbf{X} \mathbf{w}$$

What about the errors?

$$t = w_0 + w_1x = \mathbf{w}^T \mathbf{x}$$

$$t = w_0 + w_1x_1 + w_2x_2 + w_3x_3 + \dots + w_Dx_D = \sum_{d=0}^D w_dx_d = \mathbf{w}^T \mathbf{x}$$

$$\hat{\mathbf{w}} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{t}$$

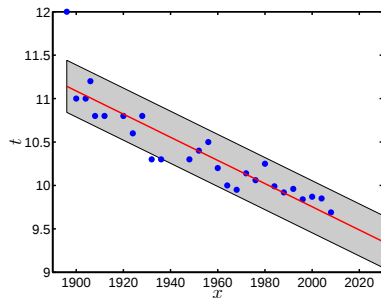


We **should** model the errors

- ▶ We know they're there - shouldn't ignore them.

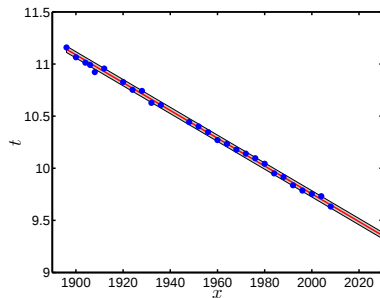
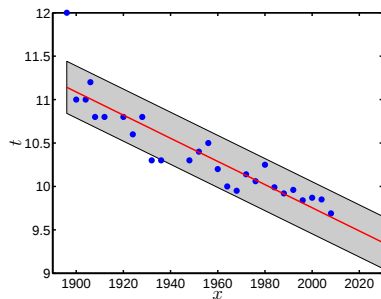
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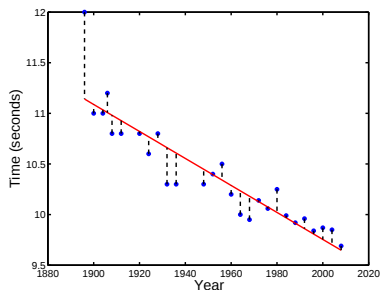


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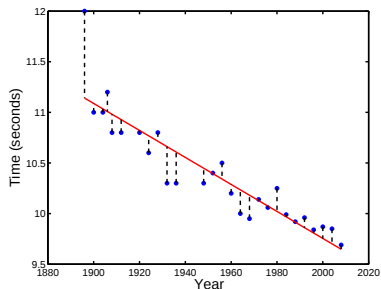
Additive errors



We'll assume that the noise is an additive term in the model:

$$t_n = \mathbf{w}^T \mathbf{x} + \epsilon_n$$

Additive errors

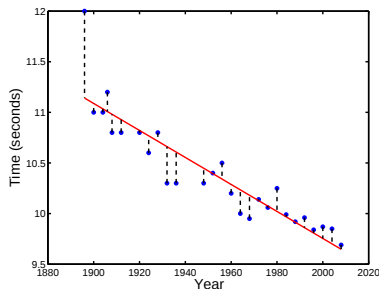


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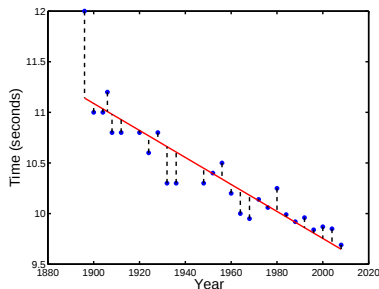
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- It's different for each n .

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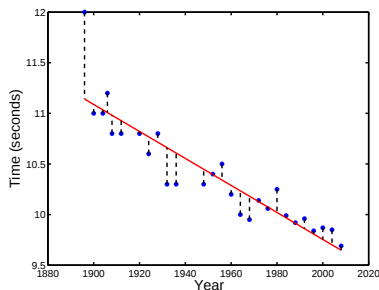
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- ▶ It's different for each n .
- ▶ It's positive and negative.

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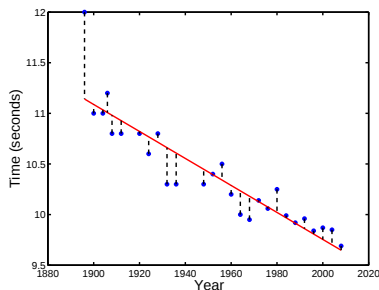
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- ▶ There doesn't seem to be any relationship between ϵ at different n .

Additive errors



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What assumptions can we make about ϵ_n ?

- ▶ It's different for each n .
- ▶ It's positive and negative.
- ▶ There doesn't seem to be any relationship between ϵ at different n .
- ▶ Looks very hard to model exactly (if it were, it wouldn't be noise!)

Gaussian noise model

- Our model:

$$t_n = \mathbf{w}^T \mathbf{x} + \epsilon_n$$

Gaussian noise model

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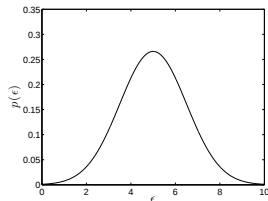
- ▶ ϵ_n is continuous.
- ▶ We need to choose $p(\epsilon)$.

Gaussian noise model

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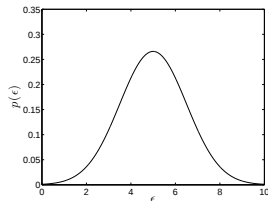
$$p(\epsilon|\mu, \sigma^2) = \frac{1}{\sigma\sqrt{2\pi}} \exp \left\{ -\frac{1}{2\sigma^2} (\epsilon - \mu)^2 \right\}$$

Gaussian noise model

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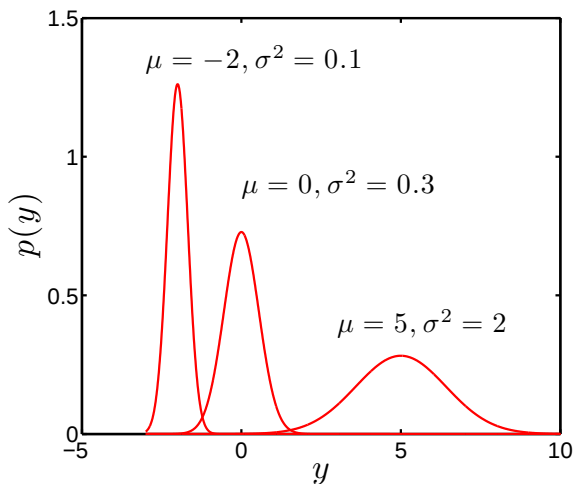
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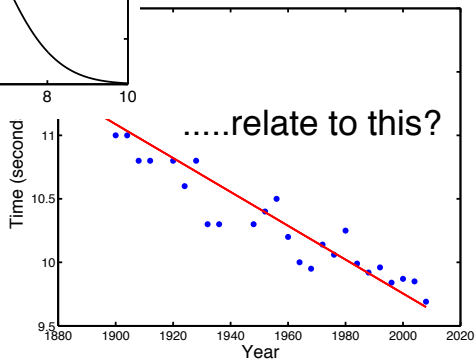
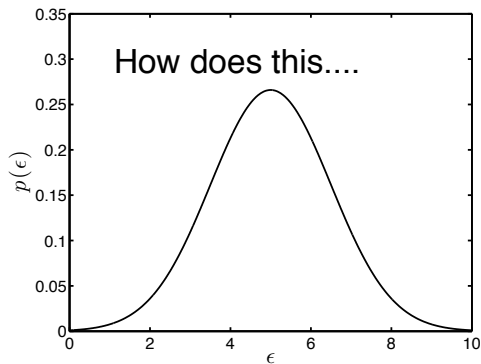
- ▶ 2 parameters: Mean μ and Variance σ^2 .

Gaussian examples



Effect of varying the mean (μ) and variance (σ^2) parameters of the Gaussian.

Generating data



Likelihood

- ▶ Evaluate the density:

$$p(t|\mathbf{x}_n, \mathbf{w}, \sigma^2) = \mathcal{N}(\mathbf{w}^\top \mathbf{x}_n, \sigma^2)$$

- ▶ t is a random variable too!
- ▶ at $t = t_n$ is called for the **Likelihood**, i.e., the quantity obtained when evaluating the density.

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- ▶ The higher the value, the more likely t_n is given the model....

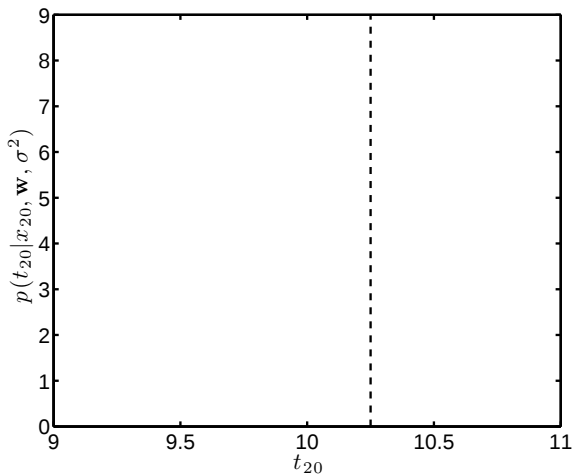
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- ▶ at $t = t_n$ is called for the **Likelihood**, i.e., the quantity obtained when evaluating the density.
- ▶ The higher the value, the more likely t_n is given the model....
 - ▶the better the model is.

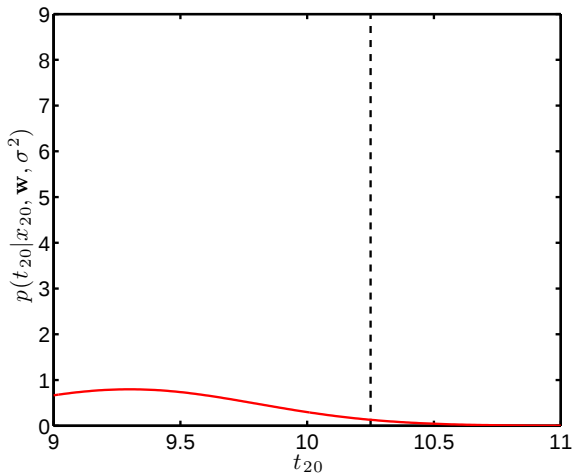
Likelihood



Lets look at the 1980 Olympics ($n = 20$).

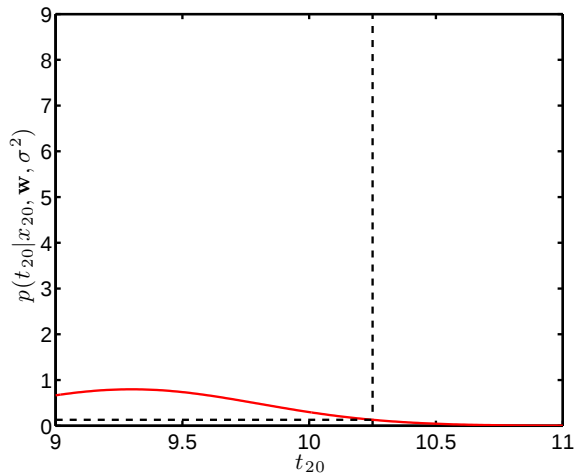
Dashed line shows t_{20} .

Likelihood



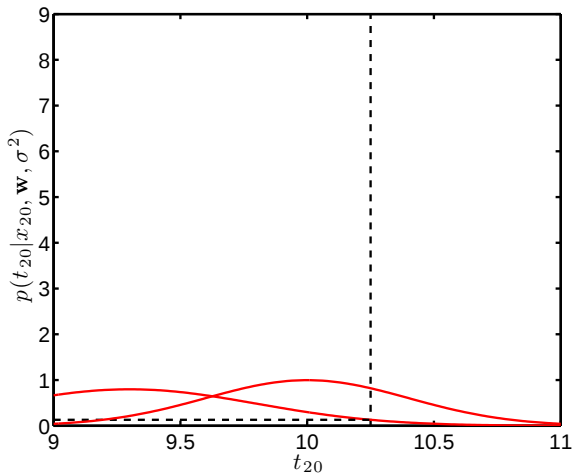
Model 1. Red line shows $\mathcal{N}(\mathbf{w}^T \mathbf{x}_n, \sigma^2)$

Likelihood



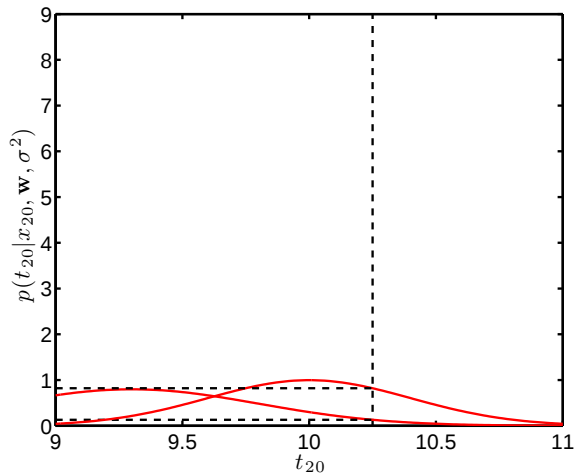
$$p(t_{20}|\dots) \approx 0.1.$$

Likelihood



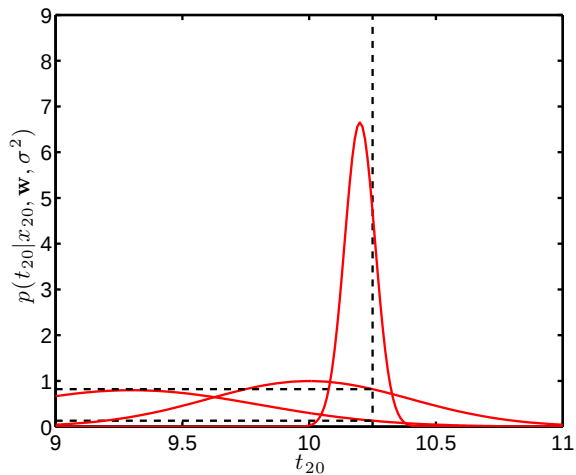
Model 2. Red line shows $\mathcal{N}(\mathbf{w}^\top \mathbf{x}_n, \sigma^2)$ for a different $\mathbf{w}$

Likelihood



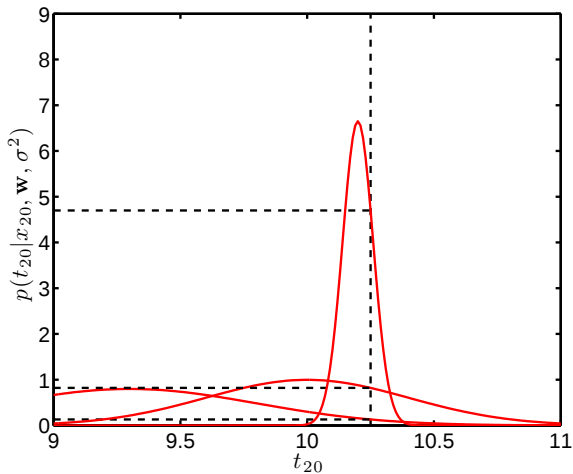
$$p(t_{20}|\dots) \approx 0.9.$$

Likelihood



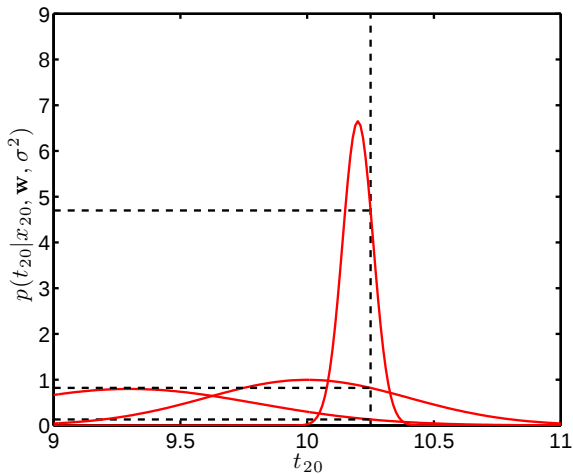
Model 3.

Likelihood



Model 3.

Likelihood



Model 3.

Model 3 looks best.

Likelihood

- ▶ The value we get when we evaluate the density function is called the **likelihood**.

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- ▶ i.e.
 - ▶ The likelihood for model 1 was 0.1.
 - ▶ The likelihood for model 2 was 0.9.
 - ▶ The likelihood for model 3 was 4.8.
- ▶ For continuous random variables, it is **not** a probability!

Likelihood

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 - ▶ The likelihood for model 1 was 0.1.
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 - ▶ The likelihood for model 3 was 4.8.
- ▶ For continuous random variables, it is **not** a probability!
- ▶ As t_n is fixed, we can find the values of $\mathbf{w}$ and σ^2 that maximise the likelihood.
 - ▶ ...just like we found them that minimised the loss.

Likelihood optimisation

- ▶ For each input-response pair, we have a Gaussian likelihood:

$$p(t_n|\mathbf{w}, \mathbf{x}_n, \sigma^2) = \mathcal{N}(\mathbf{w}^\top \mathbf{x}_n, \sigma^2)$$

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- ▶ To combine them all, we want the joint likelihood:

$$p(t_1, \dots, t_N | \mathbf{w}, \sigma^2, \mathbf{x}_1, \dots, \mathbf{x}_N)$$

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- ▶ To combine them all, we want the joint likelihood:

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- ▶ Assume that the t_n are independent:

$$p(t_1, \dots, t_N | \mathbf{w}, \sigma^2, \mathbf{x}_1, \dots, \mathbf{x}_N) = \prod_{n=1}^N p(t_n | \mathbf{w}, \mathbf{x}_n, \sigma^2)$$

Likelihood optimisation

Finding the parameters that maximise the likelihood is expressed mathematically as:

$$\operatorname{argmax}_{\mathbf{w}, \sigma^2} \prod_{n=1}^N p(t_n | \mathbf{w}, \mathbf{x}_n, \sigma^2)$$

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In fact, we'll optimise the (natural) log likelihood because it's easier.

- ▶ If we increase z , $\log(z)$ increases, if we decrease z , $\log(z)$ decreases. So, at a maximum of z , $\log(z)$ will also be at a maximum.

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$$\operatorname{argmax}_{\mathbf{w}, \sigma^2} \log \prod_{n=1}^N p(t_n | \mathbf{w}, \mathbf{x}_n, \sigma^2)$$

Some re-arranging...

$$p(t_n|\mathbf{w}, \mathbf{x}_n, \sigma^2) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left\{-\frac{1}{2\sigma^2}(t_n - \mathbf{w}^\top \mathbf{x}_n)^2\right\}$$

$$\log L = \log \prod_{n=1}^N p(t_n|\mathbf{w}, \mathbf{x}_n, \sigma^2)$$

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Looks familiar!

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Looks familiar! To continue (good exercise):

$$\frac{\partial \log L}{\partial \mathbf{w}} = 0, \quad \frac{\partial \log L}{\partial \sigma^2} = 0$$

A shortcut

The multi-variate Gaussian

$$\mathbf{y} = \begin{bmatrix} y_1 \\ y_2 \end{bmatrix}, \quad p(\mathbf{y}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$$

$$\mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{(2\pi)^{K/2} |\boldsymbol{\Sigma}|^{1/2}} \exp \left\{ -\frac{1}{2} (\mathbf{y} - \boldsymbol{\mu})^\top \boldsymbol{\Sigma}^{-1} (\mathbf{y} - \boldsymbol{\mu}) \right\}$$

$K(= 2)$ is number of variables, $|\boldsymbol{\Sigma}|$ is the determinant.

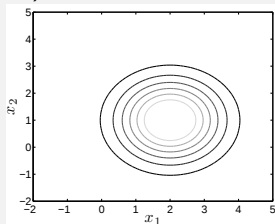
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$$\boldsymbol{\mu} = \begin{bmatrix} 2 \\ 1 \end{bmatrix}, \quad \boldsymbol{\Sigma} = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

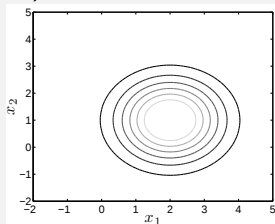
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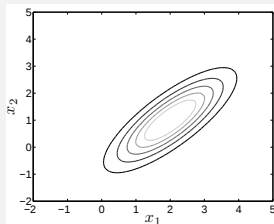
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$$\boldsymbol{\mu} = \begin{bmatrix} 2 \\ 1 \end{bmatrix}, \quad \boldsymbol{\Sigma} = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$



$$\boldsymbol{\Sigma} = \begin{bmatrix} 1 & 0.8 \\ 0.8 & 1 \end{bmatrix}$$

A shortcut

The multi-variate Gaussian

A special case:

$$\prod_{n=1}^N \mathcal{N}(\mu_n, \sigma^2) = \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$$
$$\boldsymbol{\mu} = \begin{bmatrix} \mu_1 \\ \vdots \\ \mu_N \end{bmatrix}, \quad \boldsymbol{\Sigma} = \begin{bmatrix} \sigma^2 & \dots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \dots & \sigma^2 \end{bmatrix} = \sigma^2 \mathbf{I}$$

So, in our model:

$$\log L = \log \prod_{n=1}^N p(t_n | \mathbf{w}, \mathbf{x}_n, \sigma^2) = \log \mathcal{N}(\mathbf{X}\mathbf{w}, \sigma^2 \mathbf{I}) = \log p(\mathbf{t} | \mathbf{w}, \mathbf{X}, \sigma^2)$$

Maximising the multi-variate log-likelihood

- Partial derivative w.r.t. $\mathbf{w}$, set to zero and solve:

$$\begin{aligned}\log L &= \log \mathcal{N}(\mathbf{X}\mathbf{w}, \sigma^2 \mathbf{I}) \\ \frac{\partial \log L}{\partial \mathbf{w}} &= -\frac{1}{2\sigma^2} (2\mathbf{X}^\top \mathbf{X} \mathbf{w} - 2\mathbf{X}^\top \mathbf{t}) = 0 \\ \mathbf{w} &= (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{t}\end{aligned}$$

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- ▶ This is the same expression we've seen before!

Maximising the multi-variate log-likelihood

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- ▶ This is the same expression we've seen before!
- ▶ Same for σ^2 :

$$\begin{aligned}\frac{\partial \log L}{\partial \sigma^2} &= -\frac{N}{2\sigma^2} + \frac{1}{2(\sigma^2)^2} (\mathbf{t} - \mathbf{X}\mathbf{w})^\top (\mathbf{t} - \mathbf{X}\mathbf{w}) = 0 \\ \sigma^2 &= \frac{1}{N} (\mathbf{t} - \mathbf{X}\mathbf{w})^\top (\mathbf{t} - \mathbf{X}\mathbf{w})\end{aligned}$$

Optimum parameters

- ▶ Compute optimum $\hat{\mathbf{w}}$ from:

$$\hat{\mathbf{w}} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{t}$$

- ▶ Use this to compute optimum $\hat{\sigma}^2$ from:

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Optimum parameters

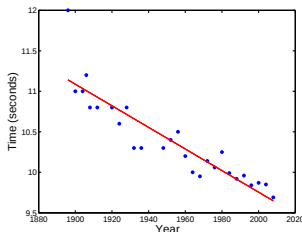
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- e.g. Olympic 100 m data (again!)



$$\hat{\mathbf{w}} = \begin{bmatrix} 36.416 \\ -0.0133 \end{bmatrix}, \quad \hat{\sigma}^2 = 0.0503$$

Optimum parameters

- ▶ We have point estimates of our parameters.
- ▶ How confident should we be in them?
 - ▶ If we changed them a little bit, would the model still be good?

Confidence in parameter estimates

- Imagine there are **true** parameters, $\mathbf{w}$ and σ^2 .

Confidence in parameter estimates

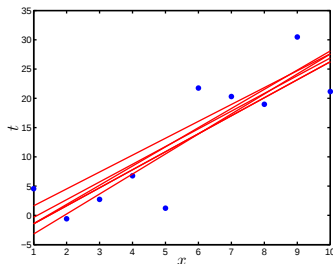
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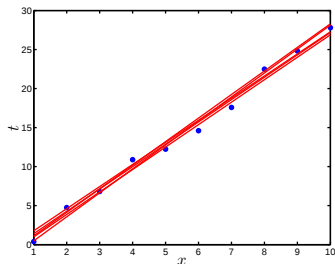
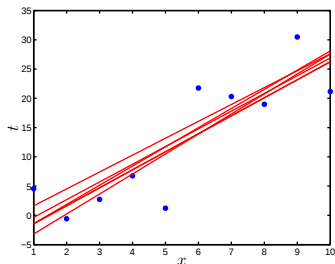
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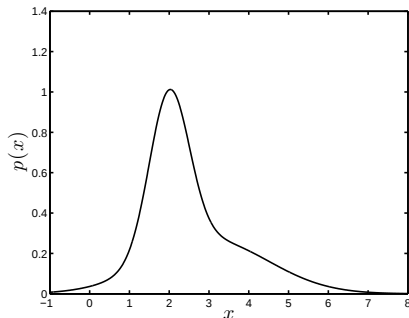


Expectations – refresher

- ▶ To progress we need to understand **Expectations**

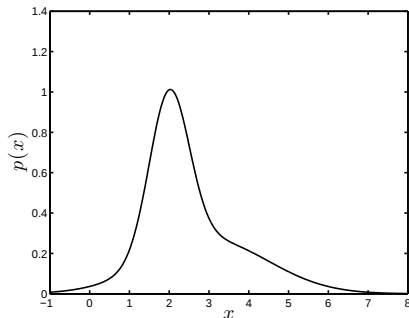
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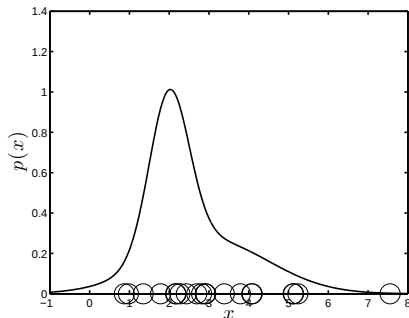
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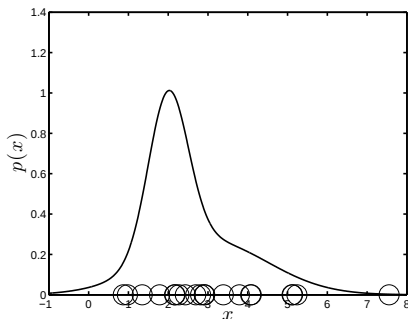
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$$\tilde{x} \approx \frac{1}{S} \sum_{s=1}^S x_s$$



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Expectations – refresher

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Expectations – Gaussians

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► Multi-variate

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Back to the model...

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- ▶ What do we expect our parameter estimate to be?

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We'll try and find $\mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \{ \hat{\mathbf{w}} \}$ in terms of the true value $\mathbf{w}$:

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$\hat{\mathbf{w}}$ is unbiased

On average, we expect our estimate to equal the true value!

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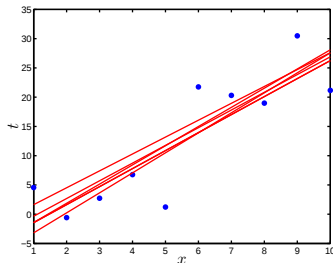
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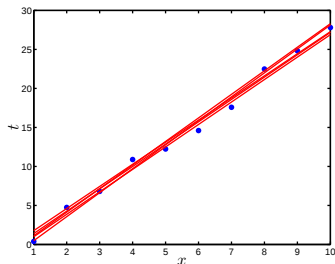
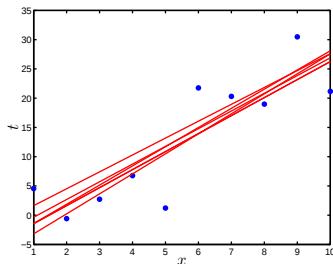


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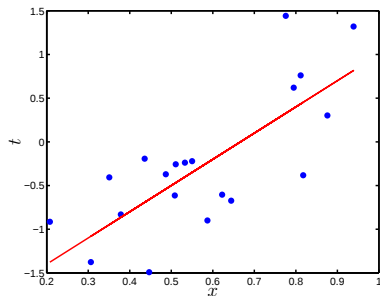
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Example

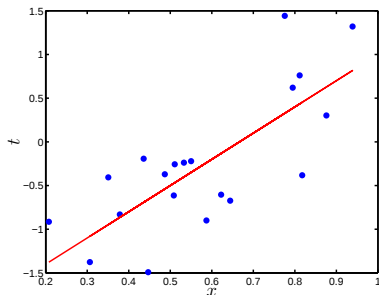


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$$p(\epsilon_n) = \mathcal{N}(0, \sigma^2)$$

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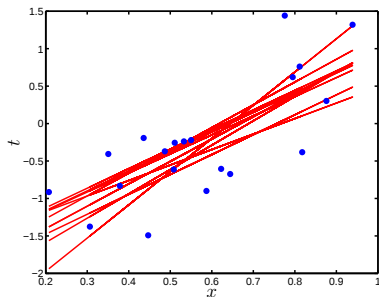
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$$\hat{\mathbf{w}} = \begin{bmatrix} -1.95 \\ 2.94 \end{bmatrix}, \quad \text{cov}\{\hat{\mathbf{w}}\} = \begin{bmatrix} 0.1195 & -0.1847 \\ -0.1847 & 0.3190 \end{bmatrix}$$

Example



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$$p(\epsilon_n) = \mathcal{N}(0, \sigma^2)$$

$$\sigma^2 = 0.5^2$$

$$\hat{\mathbf{w}} = \begin{bmatrix} -1.95 \\ 2.94 \end{bmatrix}, \quad \text{cov}\{\hat{\mathbf{w}}\} = \begin{bmatrix} 0.1195 & -0.1847 \\ -0.1847 & 0.3190 \end{bmatrix}$$

$\mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \left\{ \widehat{\sigma^2} \right\}$ – beyond this class

We saw that $\widehat{\mathbf{w}}$ was unbiased, what about $\widehat{\sigma^2}$?

$$\begin{aligned}\mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \left\{ \widehat{\sigma^2} \right\} &= \frac{1}{N} \mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \left\{ (\mathbf{t} - \mathbf{X}\widehat{\mathbf{w}})^\top (\mathbf{t} - \mathbf{X}\widehat{\mathbf{w}}) \right\} \\ &= \sigma^2 \left(1 - \frac{D}{N} \right).\end{aligned}$$

Useful identity

$$\begin{aligned}p(\mathbf{t}) &= \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \\ \mathbf{E}_{p(\mathbf{t})} \left\{ \mathbf{t}^\top \mathbf{A} \mathbf{t} \right\} &= \text{Tr}(\mathbf{A}\boldsymbol{\Sigma}) + \boldsymbol{\mu}^\top \mathbf{A} \boldsymbol{\mu} \\ \text{Tr}(\mathbf{A}) &= \sum_i A_{ii}\end{aligned}$$

Another useful identity

$$\text{Tr}(\mathbf{A}\mathbf{B}) = \text{Tr}(\mathbf{B}\mathbf{A})$$

$\mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \left\{ \widehat{\sigma^2} \right\}$ – beyond this class

$$\begin{aligned}\mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \left\{ \widehat{\sigma^2} \right\} &= \frac{1}{N} (\text{Tr}(\sigma^2 \mathbf{I}) + \mathbf{w}^\top \mathbf{X}^\top \mathbf{X} \mathbf{w}) \\ &\quad - \frac{1}{N} (\text{Tr}(\sigma^2 \mathbf{X}(\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top) + \mathbf{w}^\top \mathbf{X}^\top \mathbf{X} \mathbf{w}) \\ &= \sigma^2 - \frac{\sigma^2}{N} \text{Tr}(\mathbf{X}(\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top) \\ &= \sigma^2 - \frac{\sigma^2}{N} \text{Tr}(\mathbf{X}^\top \mathbf{X}(\mathbf{X}^\top \mathbf{X})^{-1}) \\ &= \sigma^2 \left(1 - \frac{D}{N} \right)\end{aligned}$$

Where D is the number of columns in $\mathbf{X}$ (the number of elements in $\mathbf{w}$).

Another useful identity

$$\text{Tr}(\mathbf{AB}) = \text{Tr}(\mathbf{BA})$$

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- ▶ To think about – what if $D = N$ or $D > N$?

Example – beyond this class

Generate 100 datasets from the following model:

$$t_n = w_0 + w_1 x_n + \epsilon_n, \quad p(\epsilon_n) = \mathcal{N}(0, 0.25)$$

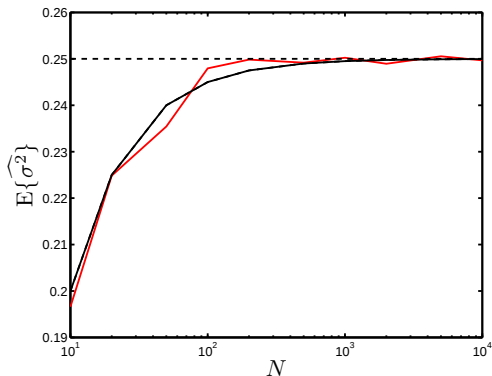
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Red curve – average $\hat{\sigma}^2$ over 100 datasets. Black curve – theoretical value. Dashed line – true value.

Summary

- ▶ Computed $\mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \{\hat{\mathbf{w}}\} = \mathbf{w}$
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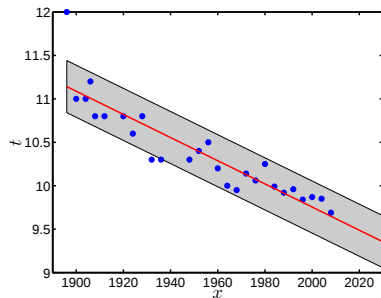
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 - ▶ $\widehat{\sigma^2}$ is **biased**.
 - ▶ Gets better and better as we get more data.

Predictions

- ▶ Our aim is to make predictions (e.g. London 2012)

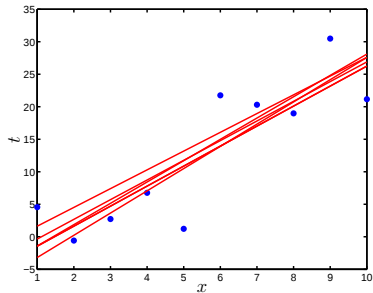
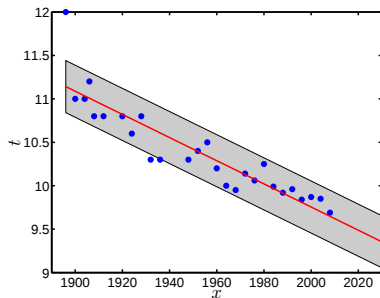
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$$t = \mathbf{w}^T \mathbf{x} + \epsilon$$

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- ▶ Is this sensible? What is $\mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \{t_{\text{new}}\}$?

$$\mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \{t_{\text{new}}\} = \mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \left\{ \hat{\mathbf{w}}^T \mathbf{x}_{\text{new}} \right\} = \mathbf{w}^T \mathbf{x}_{\text{new}}$$

- ▶ which is a good thing!

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$$\text{var}\{t_{\text{new}}\} = \mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \{t_{\text{new}}^2\} - \mathbf{E}_{p(\mathbf{t}|\mathbf{X},\mathbf{w},\sigma^2)} \{t_{\text{new}}\}^2$$

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- ▶ If the variance in the parameters is high, so is the variance in the predictions.

Example

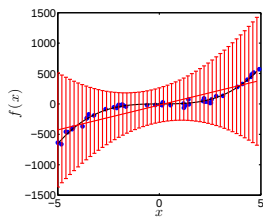
Data sampled from a 3rd order polynomial function:

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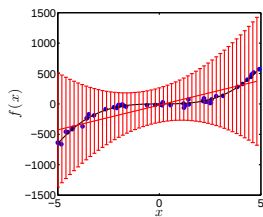
Linear

Plots show $t_{\text{new}} \pm \text{var}\{t_{\text{new}}\}$. (Black line is truth).

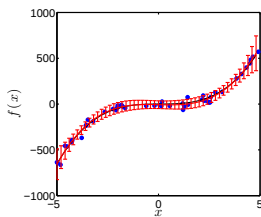
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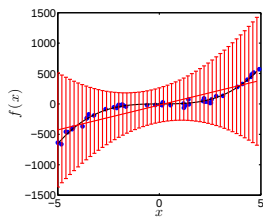
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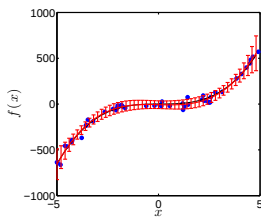
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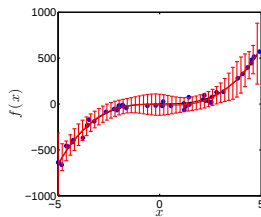
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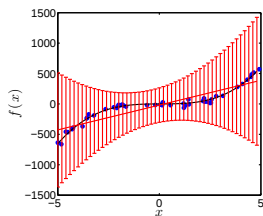
6th order

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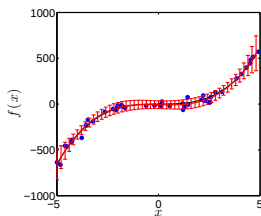
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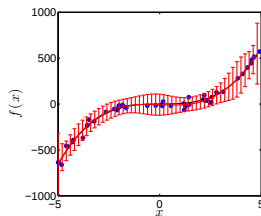
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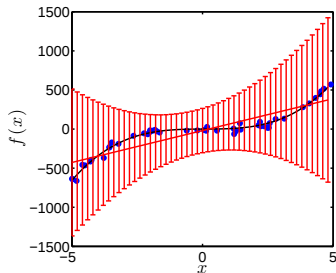
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Why does the predictive variance increase above and below the correct order?

Not complex enough model – more ‘noise’

In practice we don't know σ^2 so substitute $\widehat{\sigma^2}$:

$$\text{var}\{t_{\text{new}}\} = \widehat{\sigma^2} \mathbf{x}_{\text{new}}^T (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{x}_{\text{new}}$$

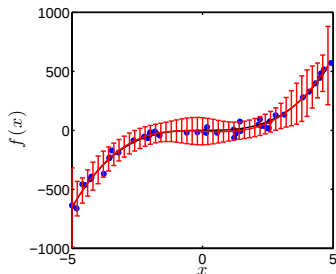


- ▶ The model is too simple.
- ▶ Some true variability can only be modelled noise.
- ▶ $\widehat{\sigma^2}$ is significantly over-estimated.
- ▶ Results in high $\text{var}\{t_{\text{new}}\}$.

Too complex model – parameters not well defined

Similarly, we substitute $\hat{\sigma}^2$ into expression for $\text{cov}\{\hat{\mathbf{w}}\}$:

$$\text{cov}\{\hat{\mathbf{w}}\} = \hat{\sigma}^2(\mathbf{X}^T\mathbf{X})^{-1}$$

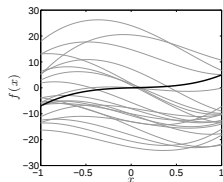
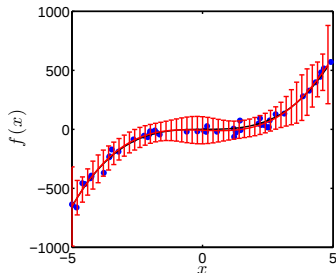


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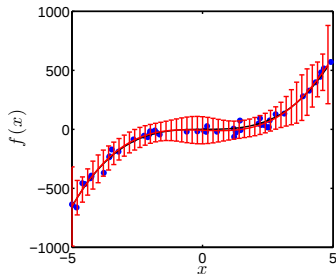


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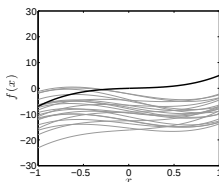
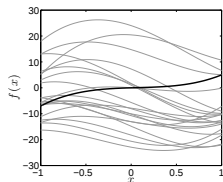
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Olympic prediction

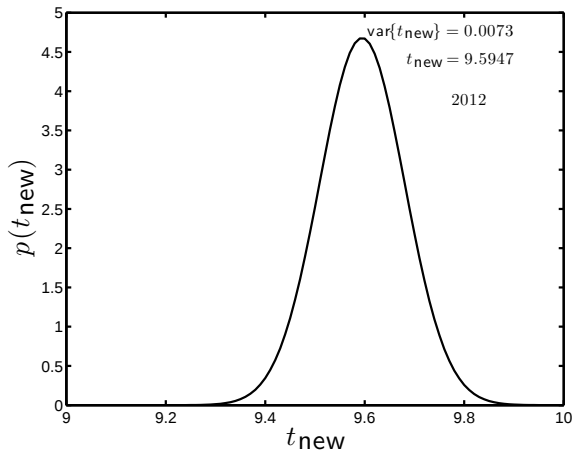
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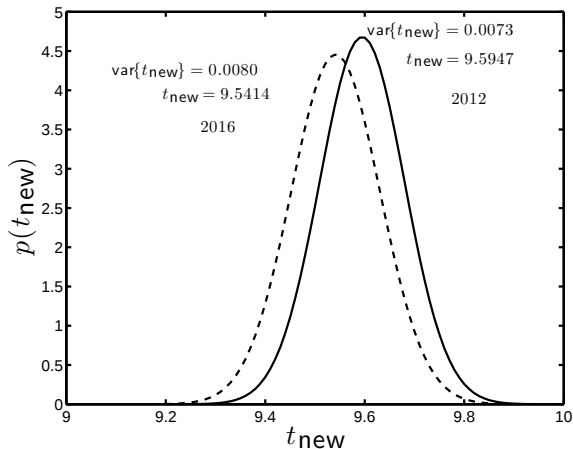
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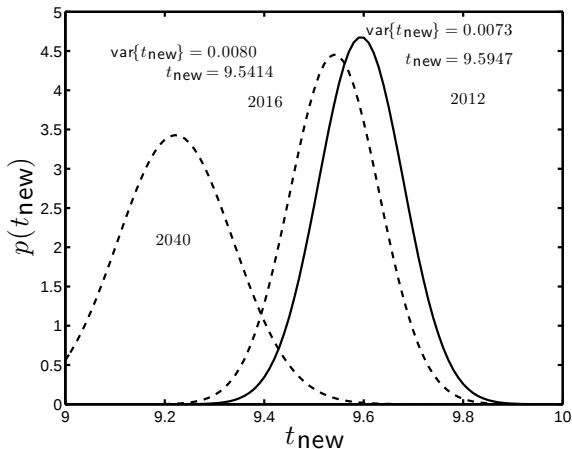


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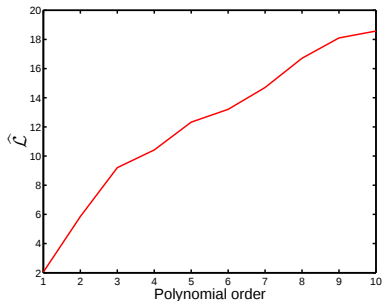
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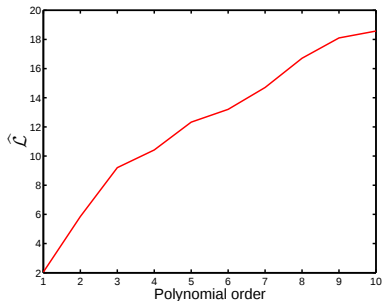


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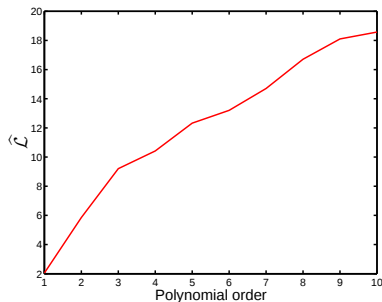


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Data from 3rd order polynomial.

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 - ▶ More complex models can always get closer to the data.
 - ▶ Results in lower $\hat{\sigma}^2$ and higher likelihood.

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- ▶ Decided to model the noise.
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- ▶ Introduced likelihood and maximised it to find $\hat{\mathbf{w}}$ and $\hat{\sigma}^2$.
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- ▶ What next?
 - ▶ Going Bayesian.
 - ▶ Got to forget about single parameter values - parameters are random variables too.

Aside - from one model to many

- ▶ All of our efforts so far have been to find the ‘best’ model:
 - ▶ The one that minimises the loss.
 - ▶ The one that maximises the likelihood.
- ▶ Given the uncertainty, maybe we shouldn’t trust one on its own?
- ▶ Consider the following random variable (RV):

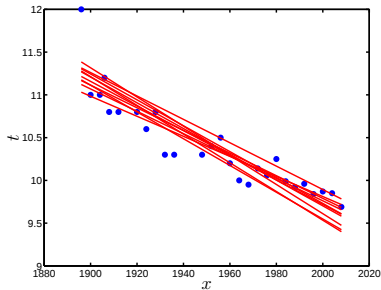
$$p(\mathbf{q}) = \mathcal{N}(\widehat{\mathbf{w}}, \text{cov}\{\widehat{\mathbf{w}}\})$$

- ▶ Samples of this RV $\mathbf{q}_s$ are **models** (assume $\widehat{\sigma}^2$ is fixed)
- ▶ We can generate lots of good models...

- ▶ Sample lots of $\mathbf{q}$ from:

$$p(\mathbf{q}) = \mathcal{N}(\widehat{\mathbf{w}}, \text{cov}\{\widehat{\mathbf{w}}\})$$

- ▶ Each corresponds to a model.

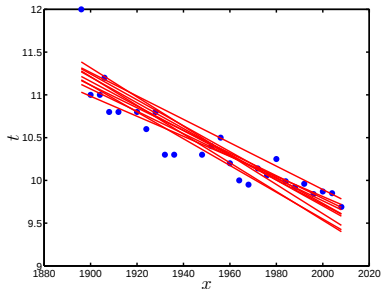


- ▶ Sample lots of $\mathbf{q}$ from:

$$p(\mathbf{q}) = \mathcal{N}(\widehat{\mathbf{w}}, \text{cov}\{\widehat{\mathbf{w}}\})$$

- ▶ Each corresponds to a model.
- ▶ Compute a prediction from each one:

$$t_s = \mathbf{q}_s^T \mathbf{x}_{\text{new}}$$



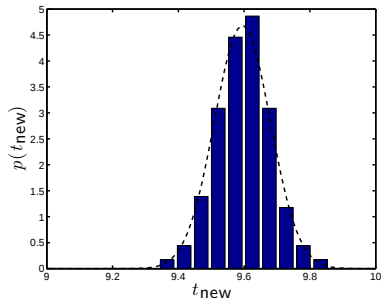
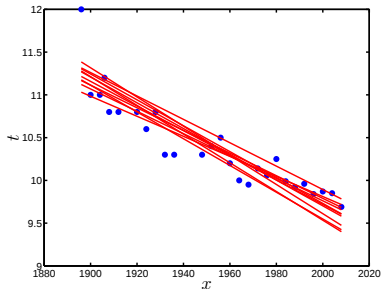
- ▶ Sample lots of $\mathbf{q}$ from:

$$p(\mathbf{q}) = \mathcal{N}(\hat{\mathbf{w}}, \text{cov}\{\hat{\mathbf{w}}\})$$

- ▶ Each corresponds to a model.
- ▶ Compute a prediction from each one:

$$t_s = \mathbf{q}_s^T \mathbf{x}_{\text{new}}$$

- ▶ Look at the distribution of predictions:



Do we need to take samples at all?

- ▶ Take an expectation...

$$\mathbf{E}_{p(\mathbf{q})} \{t_{\text{new}}\} = \int t_{\text{new}} \mathcal{N}(\hat{\mathbf{w}}, \text{cov}\{\hat{\mathbf{w}}\}) dt_{\text{new}}$$

- ▶ We'll see more of this in the next lecture....