

MVE550 2019 Lecture 10

Petter Mostad

Chalmers University

December 5, 2019

Where are we?

- ▶ In the beginning of the course, we defined a stochastic process as a collection $\{X_t, t \in I\}$ of random variables with a common state space S .
- ▶ So far, the set I has been the non-negative integers. We now move on to processes where I is a non-countable set, for example all positive real numbers, or all subsets of \mathbb{R}^2 .
- ▶ Chapters 6 and 7 of Dobrow concern such stochastic processes where the state space S is discrete.
- ▶ In Chapter 8 of Dobrow we look at the situation when the random variables X_t are continuous variables.

- ▶ Three equivalent definitions of a Poisson process.
- ▶ A number of important and useful properties.
- ▶ Examples and example computations!
- ▶ Spatial Poisson processes and inhomogeneous Poisson processes.

Counting processes

- ▶ A *counting process* $\{N_t, t \in I\}$ is a stochastic process where $I = \mathbb{R}_0^+$, where the state space is the non-negative integers, and where $0 \leq s \leq t$ implies $N_s \leq N_t$.
- ▶ Informally, when $s < t$, $N_t - N_s$ counts the number of “events” in $(s, t]$.
- ▶ N_t is a function of t that is a right-continuous step function.

Poisson process: Definiton 1

- ▶ A Poisson process $\{N_t\}_{t \geq 0}$ with parameter $\lambda > 0$ is a counting process fulfilling
 - ▶ $N_0 = 0$.
 - ▶ $N_t \sim \text{Poisson}(\lambda t)$ for all $t > 0$.
 - ▶ *Stationary increments*: $N_{t+s} - N_s$ has the same distribution as N_t .
 - ▶ *Independent increments*: $N_t - N_s$ and $N_r - N_q$ are independent, when $0 \leq q < r \leq s < t$.
- ▶ Note: Not obvious that such a process exists.
- ▶ Note: $E(N_t) = \lambda t$. Thus what one is counting occurs with a *rate* of λ items per time unit.

Poisson process: Definition 2

- ▶ Let X_1, X_2, \dots , be a sequence of iid exponential random variables with parameter λ . Define $N_0 = 0$ and, for $t > 0$,

$$N_t = \max\{n : X_1 + \dots + X_n \leq t\}.$$

Then $\{N_t\}_{t \geq 0}$ is a Poisson process with parameter λ .

- ▶ We call $S_n = X_1 + \dots + X_n$ the *arrival times* of the process.
- ▶ We call $X_k = S_k - S_{k-1}$ the *inter-arrival times* of the process.
- ▶ This provides an easy way to simulate a Poisson process.

Memorylessness of the exponential distribution

- ▶ A random variable X is called *memoryless* if

$$P(X > s + t \mid X > s) = P(X > t)$$

for all $s > 0, t > 0$.

- ▶ The exponential distribution is memoryless, and is the only memoryless continuous random variable.
- ▶ Consider the consequences of this when using the exponential as a model.

Minimum and sum of independent exponentially distributed variables

- ▶ Define $M = \min(X_1, \dots, X_n)$ where, independently for each i , $X_i \sim \text{Exponential}(\lambda_i)$. Then:
 - ▶ $M \sim \text{Exponential}(\lambda_1 + \dots + \lambda_n)$.
 - ▶ $P(M = X_k) = \frac{\lambda_k}{\lambda_1 + \dots + \lambda_n}$.
- ▶ Let $S_n = X_1 + \dots + X_n$ where, independently for each i , $X_i \sim \text{Exponential}(\lambda)$. Then $S_n \sim \text{Gamma}(n, \lambda)$.

Poisson process: Definition 3

- ▶ Introduce/review the $o(h)$ and $o(g(h))$ notation.
- ▶ A Poisson process $\{N_t\}_{t \geq 0}$ with parameter λ is a counting process fulfilling
 - ▶ $N_0 = 0$.
 - ▶ The process has stationary and independent increments.
 - ▶ We have

$$P(N_h = 0) = 1 - \lambda h + o(h)$$

$$P(N_h = 1) = \lambda h + o(h)$$

$$P(N_h > 1) = o(h)$$

- ▶ All the three definitions of a Poisson process are equivalent.

Thinned poisson processes

Let $\{N_t\}_{t \geq 0}$ be a Poisson process with parameter λ . Assume each arrival is “marked” as “type k ”, for one of n types, with probability p_k , where $p_1 + \dots + p_n = 1$. Let $N_t^{(k)}$ be the count of the number of arrivals of type k by time t . Then

- ▶ $\{N_t^{(k)}\}_{t \geq 0}$ is a Poisson process with parameter $p_k \lambda$.
- ▶ The processes

$$\{N_t^{(1)}\}_{t \geq 0}, \dots, \{N_t^{(n)}\}_{t \geq 0}$$

are independent.

Superposition process

Assume

$$\left\{N_t^{(1)}\right\}_{t \geq 0}, \dots, \left\{N_t^{(n)}\right\}_{t \geq 0}$$

are independent Poisson processes with parameters $\lambda_1, \dots, \lambda_n$, respectively. Define, for $t > 0$,

$$N_t = N_t^{(1)} + \dots + N_t^{(n)}.$$

Then $\{N_t\}_{t \geq 0}$ is a Poisson process with parameter $\lambda = \lambda_1 + \dots + \lambda_n$.

Uniform distribution when count is fixed

Let S_1, S_2, \dots , be the arrival times of a Poisson process with parameter λ . **Conditional on $N_t = n$** , we have

- ▶ The joint density function for S_1, \dots, S_n is uniform on the set $0 < s_1 < s_2 < \dots, < s_n < t$.
- ▶ Equivalently, if U_1, \dots, U_n are iid uniform on $[0, t]$, and if $U_{(1)}, \dots, U_{(n)}$ is the ordering of these random variables, then (S_1, \dots, S_n) and $(U_{(1)}, \dots, U_{(n)})$ have the same distribution.
- ▶ The upshot: If we want to simulate a Poisson process on an interval $[0, t]$, we may first simulate N_t (the total number of “events”) and then independently simulate the arrival times of each of the N_t events uniformly on $[0, t]$.

Spatial Poisson processes

- ▶ A collection of random variables $\{N_A\}_{A \subseteq \mathbb{R}^d}$ is a spatial Poisson process with parameter λ if
 - ▶ For each bounded set $A \subseteq \mathbb{R}^d$, N_A has a Poisson distribution with parameter $\lambda|A|$.
 - ▶ Whenever $A \subseteq B$, $N_A \leq N_B$.
 - ▶ Whenever A and B are disjoint sets, N_A and N_B are independent.
- ▶ How to simulate
- ▶ One may use simulations to estimate properties such as the average distance to the nearest neighbour (or the third nearest neighbour or whatever).
- ▶ Very useful model in practice.

Non-homogeneous Poisson processes

- ▶ A counting process $\{N_t\}_{t \geq 0}$ is a *non-homogeneous* Poisson process with intensity function $\lambda(t)$ if
 - ▶ $N_0 = 0$.
 - ▶ For $t > 0$,

$$N_t \sim \text{Poisson} \left(\int_0^t \lambda(x) dx \right)$$

- ▶ It has independent increments.
- ▶ Again a very flexible and useful model in practice.
- ▶ One may have non-homogeneous spatial Poisson processes.