

**Suggested solutions for  
 MVE550 Stochastic Processes and Bayesian Inference  
 Exam April 24 2019**

1. (a) A regular transition matrix  $P$  is a transition matrix such that there is an  $n > 0$  such that  $P^n$  is a positive matrix: A positive matrix is one where all the elements are positive.
- (b) A communication class is a subset  $S$  of states such that, for all  $i, j \in S$ , there are  $n > 0$  and  $m > 0$  such that  $P_{ij}^n > 0$  and  $P_{ji}^m > 0$ , while for any pair  $i \in S$  and  $j \notin S$ , this is not the case. A closed communication class is a communication class with a zero probability of ever leaving the class.
- (c) A state  $j$  is transient if the probability that a chain starting at  $j$  will ever return to  $j$  is less than 1.
- (d) A state  $j$  is positive recurrent if the expected number of steps for a chain to return to  $j$  if it starts at  $j$  is finite.
- (e) A finite state space Markov chain is ergodic if it is irreducible and aperiodic.
- (f) Time reversibility means that, for all states  $i$  and  $j$ ,  $\pi_i P_{ij} = \pi_j P_{ji}$ .

2. (a) Using Bayes theorem we get

$$\begin{aligned}
 \pi(\lambda | x) &\propto_{\lambda} \pi(x | \lambda)\pi(\lambda) \\
 &\propto_{\lambda} \text{Exponential}(x; \lambda) \text{Gamma}(\lambda; \alpha, \beta) \\
 &\propto_{\lambda} \lambda \cdot \exp(-\lambda x) \cdot \lambda^{\alpha-1} \cdot \exp(-\beta\lambda) \\
 &\propto_{\lambda} \lambda^{\alpha} \cdot \exp(-(\beta + x)\lambda) \\
 &\propto_{\lambda} \text{Gamma}(\lambda; \alpha + 1, \beta + x)
 \end{aligned}$$

In other words, the posterior distribution is a Gamma distribution with parameters  $\alpha + 1$  and  $\beta + x$ .

- (b) The prior corresponds to a Gamma(0, 0) distribution. The posterior is obtained by updating the Gamma distribution as in (a) with the data given, resulting in the posterior

$$\text{Gamma}(3, 1.2 + 1.7 + 0.9) = \text{Gamma}(3, 3.8).$$

3. (a) We get the transition matrix

$$P = \begin{bmatrix} 0.1 & 0.3 & 0.6 & 0 \\ 0.1 & 0.2 & 0.3 & 0.4 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix}.$$

(b) We get

$$Q = \begin{bmatrix} 0.1 & 0.3 \\ 0.1 & 0.2 \end{bmatrix}$$

and so

$$\begin{aligned} F &= (I - Q)^{-1} = \begin{bmatrix} 0.9 & -0.3 \\ -0.1 & 0.8 \end{bmatrix}^{-1} = \frac{1}{0.9 \cdot 0.8 - 0.3 \cdot 0.1} \begin{bmatrix} 0.8 & 0.3 \\ 0.1 & 0.9 \end{bmatrix} \\ &= \frac{1}{0.69} \begin{bmatrix} 0.8 & 0.3 \\ 0.1 & 0.9 \end{bmatrix} = \begin{bmatrix} 1.1594 & 0.4348 \\ 0.1449 & 1.3043 \end{bmatrix}. \end{aligned}$$

(c) We have

$$R = \begin{bmatrix} 0.6 & 0 \\ 0.3 & 0.4 \end{bmatrix}$$

and thus

$$FR = \frac{1}{0.69} \begin{bmatrix} 0.8 & 0.3 \\ 0.1 & 0.9 \end{bmatrix} \begin{bmatrix} 0.6 & 0 \\ 0.3 & 0.4 \end{bmatrix} = \frac{1}{0.69} \begin{bmatrix} 0.57 & 0.12 \\ 0.33 & 0.36 \end{bmatrix}.$$

Thus the probability for a process that starts in state 1 to be absorbed in state 4 is  $\frac{0.12}{0.69} = 0.1739$ .

4. Assume  $X_0, X_1, \dots, X_n, \dots$  is an ergodic Markov chain with stationary distribution  $\pi$ . Let  $r$  be a bounded real-valued function. Then

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n r(X_i) = E(r(X))$$

where  $X$  is a random variable with distribution  $\pi$ .

5. The holding time parameters are  $q = (2, 3, 4)$ . The embedded chain transition matrix is

$$\tilde{P} = \begin{bmatrix} 0 & 0.6 & 0.4 \\ 0.9 & 0 & 0.1 \\ 1 & 0 & 0 \end{bmatrix}.$$

Thus the generator matrix becomes

$$Q = \begin{bmatrix} -2 & 1.2 & 0.8 \\ 2.7 & -3 & 0.3 \\ 4 & 0 & -4 \end{bmatrix}.$$

The linear system  $\pi Q = 0$  gives

$$\begin{aligned} -2\pi_1 + 2.7\pi_2 + 4\pi_3 &= 0 \\ 1.2\pi_1 - 3\pi_2 &= 0 \\ 0.8\pi_1 + 0.3\pi_2 - 4\pi_3 &= 0 \end{aligned}$$

with solution  $\pi = \frac{1}{163}(100, 40, 23)$ . Thus, the long-term proportion of time that the component spends in state A is  $100/163 = 0.6135$ .

6. A hidden Markov model consists of a Markov chain  $X_1, X_2, \dots, X_n$  of “hidden” random variables, and another sequence  $Y_1, \dots, Y_n$  of variables such that the distribution of  $Y_i$  only depends on  $X_i$ , and possibly on  $Y_{i-1}$ . These latter variables are the “observed” variables. If the values of the variables  $Y_i$  are indeed observed, the posterior distribution for one of the hidden variables, say  $X_i$  can be found as follows: In a “Forward” part of the algorithm, for  $j = 1, \dots, i$ , the posterior for  $X_j$  given  $Y_1, \dots, Y_j$  is found in a recursive algorithm. In a “Backward” part of the algorithm, for  $j = n, \dots, i$ , the likelihoods for  $X_j$  given the data  $Y_j, \dots, Y_n$  are found in a recursive algorithm. Then the two are put together to find the marginal posterior for  $X_i$ .

7. (a) The events that three requests of type A arrive during the time unit and that four requests of type B arrive during the time unit are independent, and the probability of both can be computed using the Poisson probability mass function. Thus the answer is

$$e^{-\lambda_A} \frac{\lambda_A^3}{3!} e^{-\lambda_B} \frac{\lambda_B^4}{4!} = e^{-3} \frac{3^3}{3!} e^{-2} \frac{2^4}{4!} = 3e^{-5} = 0.02021384.$$

(b) The probability of the first event being a request of type A or B is  $\frac{\lambda_A}{\lambda_A + \lambda_B}$ , respectively  $\frac{\lambda_B}{\lambda_A + \lambda_B}$ . As the successive events are independent, the probability asked for is

$$\left( \frac{\lambda_A}{\lambda_A + \lambda_B} \right)^3 \left( \frac{\lambda_B}{\lambda_A + \lambda_B} \right)^4 = \frac{\lambda_A^3 \lambda_B^4}{(\lambda_A + \lambda_B)^7}.$$

8. We have

$$E(X_t) = a + b(E(B_t) + E(W_{3+t})) = a + b(0 + W_3).$$

Setting this to zero gives  $a + bW_3 = 0$ . Further,

$$\text{Var}(X_t) = b^2 (\text{Var}(B_t) + \text{Var}(W_{3+t})) = b^2(t + t)$$

and setting this to  $t$  gives  $2b^2 = 1$ . Thus we must have  $b = \frac{1}{\sqrt{2}}$  and  $a = -\frac{1}{\sqrt{2}}W_3$ . On the other hand, with these values,

$$a + b(B_t + W_{3+t}) = \frac{1}{\sqrt{2}}(B_t + (W_{3+t} - W_3))$$

fulfills all criteria for a Brownian motion.