# Bayesian Regression

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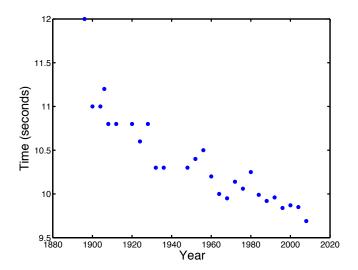
#### Reference

The content and the slides are adapted from

S. Rogers and M. Girolami, A First Course in Machine Learning (FCML), 2nd edition, Chapman & Hall/CRC 2016, ISBN: 9781498738484

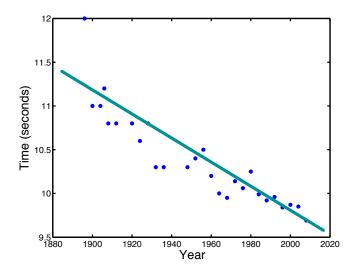
## Some data and a problem

Predict the winning time for 2012!



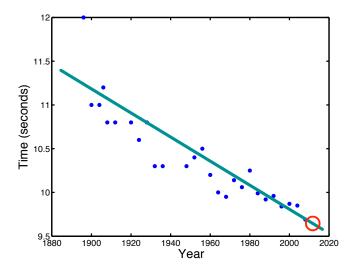
## Some data and a problem

Fit a linear model (draw a line through the data)



## Some data and a problem

Use the model (line) to *predict* the winning time in 2012.



### Recipe for a linear model

More complex model:  $t = w_0 + w_1x_1 + w_2x_2 + \ldots + w_Dx_D$ 

$$\mathbf{x}_{n} = \begin{bmatrix} 1 \\ x_{n,1} \\ x_{n,2} \\ \vdots \\ x_{n,D} \end{bmatrix}, \mathbf{X} = \begin{bmatrix} 1 & x_{1,1} & x_{1,2} & \dots & x_{1,D} \\ 1 & x_{2,1} & x_{2,2} & \dots & x_{2,D} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & x_{N,1} & x_{N,2} & \dots & x_{N,D} \end{bmatrix} \mathbf{t} = \begin{bmatrix} t_{1} \\ t_{n} \\ \vdots \\ t_{N} \end{bmatrix},$$

## Recipe for a linear model

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$$\mathbf{w} = \begin{bmatrix} w_0 \\ w_1 \\ \vdots \\ w_D \end{bmatrix}, \quad Model: t_n = \mathbf{w}^\mathsf{T} \mathbf{x}_n, \quad or \quad \mathbf{t} = \mathbf{X} \mathbf{w}$$

## Recipe for linear model

Model: 
$$t_n = \mathbf{w}^\mathsf{T} \mathbf{x}_n$$
, or  $\mathbf{t} = \mathbf{X} \mathbf{w}$ 

Usually,  $\mathbf{t}$  and  $\mathbf{X}\mathbf{w}$  are not exactly equal. So, we try to minimise the difference.

$$\mathcal{L} = \frac{1}{N} (\mathbf{t} - \mathbf{X} \mathbf{w})^{\mathsf{T}} (\mathbf{t} - \mathbf{X} \mathbf{w})$$

$$\widehat{\boldsymbol{\mathsf{w}}} = (\boldsymbol{\mathsf{X}}^{\mathsf{T}}\boldsymbol{\mathsf{X}})^{-1}\boldsymbol{\mathsf{X}}^{\mathsf{T}}\boldsymbol{\mathsf{t}}$$

## Recipe for a linear model

Model

$$t_n = \mathbf{w}^\mathsf{T} \mathbf{x}_n$$
, or  $\mathbf{t} = \mathbf{X} \mathbf{w}$ 

**Parameters** 

$$\widehat{\boldsymbol{w}} = (\boldsymbol{X}^{\mathsf{T}}\boldsymbol{X})^{-1}\boldsymbol{X}^{\mathsf{T}}\boldsymbol{t}$$

Prediction

$$\mathbf{x}_{\mathsf{new}} = \left[ egin{array}{c} 1 \\ x_{\mathsf{new},1} \\ x_{\mathsf{new},2} \\ \vdots \\ x_{\mathsf{new},D} \end{array} 
ight]$$

then compute

$$t_{\text{new}} = \widehat{\mathbf{w}}^{\mathsf{T}} \mathbf{x}_{\text{new}}$$

▶ In the probabilistic linear regression, we model the error, i.e.,

Model: 
$$t_n = \mathbf{w}^\mathsf{T} \mathbf{x}_n + \epsilon_n$$
, or  $\mathbf{t} = \mathbf{X} \mathbf{w} + \epsilon$ 

In other words, we consider  $p(t_n|\mathbf{w}, \mathbf{x}_n, \sigma^2) = \mathcal{N}(\mathbf{w}^\mathsf{T}\mathbf{x}_n, \sigma^2)$ 

▶ The full likelihood is

$$p(\mathbf{t}|\mathbf{w}, \mathbf{X}, \sigma^2) = p(t_1, \dots, t_N|\mathbf{w}, \sigma^2, \mathbf{x}_1, \dots, \mathbf{x}_N)$$

Note that

$$p(t_1,\ldots,t_N|\mathbf{w},\sigma^2,\mathbf{x}_1,\ldots,\mathbf{x}_N)=\prod_{n=1}^N p(t_n|\mathbf{w},\mathbf{x}_n,\sigma^2)$$

And  $p(\mathbf{t}|\mathbf{w}, \mathbf{X}, \sigma^2) = \mathcal{N}(\mathbf{X}\mathbf{w}, \sigma^2\mathbf{I})$ I is the identity matrix of size  $N \times N$ . The covariance marix  $\sigma^2\mathbf{I}$  indicates i.i.d..



The full likelihood is

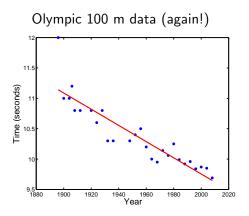
$$p(\mathbf{t}|\mathbf{w}, \mathbf{X}, \sigma^2) = p(t_1, \dots, t_N|\mathbf{w}, \sigma^2, \mathbf{x}_1, \dots, \mathbf{x}_N)$$

- We maximise the log-likelihood to obtain the parameters  ${\bf w}$  and  $\sigma^2$ .
- ► Compute optimum  $\widehat{\mathbf{w}}$  from:

$$\widehat{\mathbf{w}} = (\mathbf{X}^\mathsf{T}\mathbf{X})^{-1}\mathbf{X}^\mathsf{T}\mathbf{t}$$

• Use this to compute optimum  $\widehat{\sigma^2}$  from:

$$\widehat{\sigma^2} = \frac{1}{N} (\mathbf{t} - \mathbf{X} \widehat{\mathbf{w}})^\mathsf{T} (\mathbf{t} - \mathbf{X} \widehat{\mathbf{w}})$$



$$\widehat{\mathbf{w}} = \begin{bmatrix} 36.416 \\ -0.0133 \end{bmatrix}, \ \widehat{\sigma^2} = 0.0503$$

Model

$$p(\mathbf{t}|\mathbf{w}, \mathbf{X}, \sigma^2) = \mathcal{N}(\mathbf{X}\mathbf{w}, \sigma^2\mathbf{I})$$

**Parameters** 

$$\widehat{\boldsymbol{w}} = (\mathbf{X}^{\mathsf{T}}\mathbf{X})^{-1}\mathbf{X}^{\mathsf{T}}\mathbf{t}$$

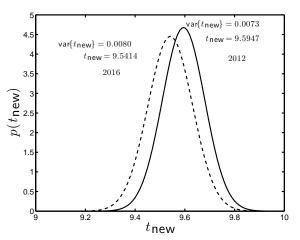
$$\widehat{\sigma^2} = \frac{1}{N}(\mathbf{t} - \mathbf{X}\widehat{\mathbf{w}})^{\mathsf{T}}(\mathbf{t} - \mathbf{X}\widehat{\mathbf{w}})$$

Prediction

$$\begin{split} t_{\mathsf{new}} &= \widehat{\mathbf{w}}^\mathsf{T} \mathbf{x}_{\mathsf{new}} \\ \mathsf{var} \{ t_{\mathsf{new}} \} &= \widehat{\sigma^2} \mathbf{x}_{\mathsf{new}}^\mathsf{T} (\mathbf{X}^\mathsf{T} \mathbf{X})^{-1} \mathbf{x}_{\mathsf{new}} \end{split}$$

*Hint:* Always check the consistency of the dimesions (numpy.shape() in Python).

### Olympic prediction



Predictive variance increases as we get further from the training data.

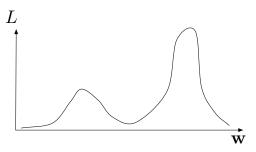
#### What is next?

- We have seen two ways of finding the 'best' parameter values:
  - ► Those that minimise the *loss L*.
  - Those that maximise the likelihood (probabilistic linear regression).
  - ▶ If the probabilistic model is Gaussian, both are the same:

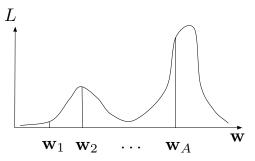
$$\widehat{\boldsymbol{\mathsf{w}}} = (\boldsymbol{\mathsf{X}}^{\mathsf{T}}\boldsymbol{\mathsf{X}})^{-1}\boldsymbol{\mathsf{X}}^{\mathsf{T}}\boldsymbol{\mathsf{t}}$$

- ▶ In the probabilistic linear regression, we also estimate  $\sigma^2$ .
- Is this the 'right' set of parameters?
- ▶ Is there a 'right' set of parameters?

### Problems with a point estimate



- ▶ Might be more than one 'best' value.
- ▶ Might not be a single representative value.
- ▶ Different values might give very different predictions.
- ► Is there an alternative?



- Prediction is some function of **w**. Say  $f(\mathbf{w})$ .
- ► Choose A different values  $-\mathbf{w}_1, \ldots, \mathbf{w}_A$ .
- ightharpoonup Compute  $\sum_{a=1}^{A} q_a f(\mathbf{w}_a)$
- $q_a$  is proportional to L (subject to  $\sum_a q_a = 1$ )
- ightharpoonup Note that each  $\mathbf{w}_a$  is a vector.
- ▶ Increasing A seems like a good idea....

## Example

- Olympic 100 m data.
- ▶ Want to predict winning time at London 2012 t<sub>new</sub>.
- ► Choose 2 'good' values of w
  - **w**<sub>1</sub> predicts  $t_{\text{new}} = 9.5 \text{ s}$
  - $\mathbf{w}_2$  predicts  $t_{\text{new}} = 9.2 \ s$
- ightharpoonup According to likelihood,  $\mathbf{w}_2$  is twice as likely as  $\mathbf{w}_1$ .

  - ▶ Therefore:  $q_1 = 1/3$ ,  $q_2 = 2/3$
- ▶ Average prediction is  $(1/3) \times 9.5 + (2/3) \times 9.2 = 9.3$

- ▶ What if **w** is a random variable with density  $p(\mathbf{w}|\text{stuff})$ ?
- Imagine a weird die that chucks out values of w.

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  - We can use every value of w!
  - We do this with the following expectation:

$$\mathbf{E}_{p(\mathbf{w}|\text{stuff})}\left\{f(\mathbf{w})\right\} = \int f(\mathbf{w})p(\mathbf{w}|\text{stuff}) \ d\mathbf{w}$$

What is  $f(\mathbf{w})$  is this course?

An average of predictions from each possible **w** weighted by how likely that **w** value is.

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- What is  $f(\mathbf{w})$  is this course?
- An average of predictions from each possible **w** weighted by how likely that **w** value is.
- ▶ What is 'stuff'?
- ▶ How do we compute  $p(\mathbf{w}|\text{stuff})$ ?

- ▶ 'Stuff' should include data: X, t: p(w|X, t)
  - ▶ i.e. what we know about **w** after observing some data.
- ▶ We've seen something like this before:  $p(\mathbf{t}|\mathbf{w}, \mathbf{X}, \sigma^2)$  the likelihood.
  - For simplicity, we ignore  $\sigma^2$  for now (we can assume its value is known).

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- Bayes rule:

$$p(\mathbf{w}|\mathbf{X},\mathbf{t}) = \frac{p(\mathbf{t}|\mathbf{X},\mathbf{w})p(\mathbf{w})}{p(\mathbf{t}|\mathbf{X})}$$

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$$p(\mathbf{w}|\mathbf{X},\mathbf{t}) = rac{p(\mathbf{t}|\mathbf{X},\mathbf{w})p(\mathbf{w})}{p(\mathbf{t}|\mathbf{X})}$$

Comes from:

$$\rho(\mathbf{w}|\mathbf{X},\mathbf{t})\rho(\mathbf{t}|\mathbf{X}) = \rho(\mathbf{t}|\mathbf{w},\mathbf{X})\rho(\mathbf{w})$$
$$\rho(\mathbf{w},\mathbf{t}|\mathbf{X}) = \rho(\mathbf{w},\mathbf{t}|\mathbf{X})$$



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- **Posterior density**:  $p(\mathbf{w}|\mathbf{X},\mathbf{t})$ 
  - ► This is what we're after.

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- **Posterior density**: p(w|X,t)
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- ▶ Prior density:  $p(\mathbf{w})$ 
  - This is new: do we know anything about the parameters before we see any data?

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  - This is what we're after.
- ▶ Likelihood :  $p(\mathbf{t}|\mathbf{X}, \mathbf{w})$ 
  - We've used this before.
- Prior density:  $p(\mathbf{w})$ 
  - This is new: do we know anything about the parameters before we see any data?
- Marginal likelihood (or evidence or normalization):
  p(t|X)
  - This is new: **w** isn't in here. It is a normalisation constant. Ensures  $\int p(\mathbf{w}|\mathbf{X},\mathbf{t}) \ d\mathbf{w} = 1$ .

## Computing the posterior

- Unfortunately, computing the posterior can be hard in general...
- ▶ ...because marginal likelihood  $p(\mathbf{t}|\mathbf{X})$  is hard to compute:

$$p(\mathbf{t}|\mathbf{X}) = \int p(\mathbf{t}|\mathbf{w}, \mathbf{X})p(\mathbf{w}) d\mathbf{w}$$

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In some cases we can do it (this lecture).

### When can we compute the posterior?

#### Conjugacy (definition)

A prior  $p(\mathbf{w})$  is said to be conjugate to a likelihood it results in a posterior of the same type of density as the prior.

- Example:
  - Prior: Gaussian; Likelihood: Gaussian; Posterior: Gaussian
  - Prior: Beta; Likelihood: Binomial; Posterior: Beta
  - ► Many others, e.g. http://en.wikipedia.org/wiki/Conjugate\_prior

# Why is this important?

$$p(\mathbf{w}|\mathbf{X},\mathbf{t}) = \frac{p(\mathbf{t}|\mathbf{X},\mathbf{w})p(\mathbf{w})}{p(\mathbf{t}|\mathbf{X})}$$

- If prior and likelihood are conjugate, we **know** the form of  $p(\mathbf{w}|\mathbf{X},\mathbf{t})$
- ▶ Therefore, we **know** the form of the normalising constant.
- ▶ Therefore, we **don't need** to compute  $p(\mathbf{t}|\mathbf{X})$

# Why is this important?

$$p(\mathbf{w}|\mathbf{X}, \mathbf{t}) = \frac{p(\mathbf{t}|\mathbf{X}, \mathbf{w})p(\mathbf{w})}{p(\mathbf{t}|\mathbf{X})}$$

- If prior and likelihood are conjugate, we **know** the form of  $p(\mathbf{w}|\mathbf{X},\mathbf{t})$
- ▶ Therefore, we **know** the form of the normalising constant.
- ▶ Therefore, we **don't need** to compute  $p(\mathbf{t}|\mathbf{X})$
- We just need to use some algebra to make  $p(\mathbf{t}|\mathbf{X}, \mathbf{w})p(\mathbf{w})$  look like the correct density, ignoring all terms without  $\mathbf{w}$ .

## Example - Olympic data

► Remember the (Gaussian) likelihood we used for maximum likelihood:

$$p(t|\mathbf{x}_n, \mathbf{w}, \sigma^2) = \mathcal{N}(\mathbf{w}^\mathsf{T}\mathbf{x}_n, \sigma^2)$$

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▶ For the set of N observations (variables)  $\{X, t\}$ , we have

$$p(\mathbf{t}|\mathbf{w}, \mathbf{X}, \sigma^2) = \mathcal{N}(\mathbf{X}\mathbf{w}, \sigma^2\mathbf{I})$$



## Example - Olympic data

► We'll use the (Gaussian) likelihood we used for maximum likelihood:

$$p(\mathbf{t}|\mathbf{w}, \mathbf{X}, \sigma^2) = \mathcal{N}(\mathbf{X}\mathbf{w}, \sigma^2\mathbf{I})$$

▶ The prior conjugate to the Gaussian is Gaussian. So:

$$p(\mathbf{w}) = \mathcal{N}(\mathbf{0}, \mathbf{S}), \ \mathbf{S} = \begin{bmatrix} 100 & 0 \\ 0 & 5 \end{bmatrix}$$

Mean (0) and covariance (S) are design choices (prior knowledge).

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- ▶ Mean (0) and covariance (S) are design choices (prior knowledge).
- **Posterior must be** Gaussian with unknown parameters  $\mu, \Sigma$ :

$$p(\mathbf{w}|\mathbf{X}, \mathbf{t}, \sigma^2) = \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$$

### Finding posterior parameters

▶ Ignoring normalising constant, the posterior is:

$$\begin{split} \rho(\mathbf{w}|\mathbf{X},\mathbf{t},\sigma^2) & \propto & \exp\left\{-\frac{1}{2}(\mathbf{w}-\boldsymbol{\mu})^\mathsf{T}\mathbf{\Sigma}^{-1}(\mathbf{w}-\boldsymbol{\mu})\right\} \\ & = & \exp\left\{-\frac{1}{2}(\mathbf{w}^\mathsf{T}\mathbf{\Sigma}^{-1}\mathbf{w}-2\mathbf{w}^\mathsf{T}\mathbf{\Sigma}^{-1}\boldsymbol{\mu}+\boldsymbol{\mu}^\mathsf{T}\mathbf{\Sigma}^{-1}\boldsymbol{\mu})\right\} \\ & \propto & \exp\left\{-\frac{1}{2}(\mathbf{w}^\mathsf{T}\mathbf{\Sigma}^{-1}\mathbf{w}-2\mathbf{w}^\mathsf{T}\mathbf{\Sigma}^{-1}\boldsymbol{\mu})\right\} \end{split}$$

▶ We only care about the terms that are related to w.

## Finding posterior parameters

▶ Ignoring non w terms, the prior multiplied by the likelihood is:

$$\begin{split} & \rho(\mathbf{t}|\mathbf{w}, \mathbf{X}, \sigma^2) \cdot \rho(\mathbf{w}) \\ & \propto & \exp\left\{-\frac{1}{2\sigma^2}(\mathbf{t} - \mathbf{X}\mathbf{w})^\mathsf{T}(\mathbf{t} - \mathbf{X}\mathbf{w})\right\} \exp\left\{-\frac{1}{2}\mathbf{w}^\mathsf{T}\mathbf{S}^{-1}\mathbf{w}\right\} \\ & \propto & \exp\left\{-\frac{1}{2}\left(\mathbf{w}^\mathsf{T}\left[\frac{1}{\sigma^2}\mathbf{X}^\mathsf{T}\mathbf{X} + \mathbf{S}^{-1}\right]\mathbf{w} - \frac{2}{\sigma^2}\mathbf{w}^\mathsf{T}\mathbf{X}^\mathsf{T}\mathbf{t}\right)\right\} \end{split}$$

Posterior (from previous slide):

$$\propto \exp\left\{-\frac{1}{2}(\mathbf{w}^\mathsf{T}\mathbf{\Sigma}^{-1}\mathbf{w} - 2\mathbf{w}^\mathsf{T}\mathbf{\Sigma}^{-1}\boldsymbol{\mu})\right\}$$

## Finding posterior parameters

- Equate individual terms on each side.
- Covariance:

$$\mathbf{w}^{\mathsf{T}} \mathbf{\Sigma}^{-1} \mathbf{w} = \mathbf{w}^{\mathsf{T}} \left[ \frac{1}{\sigma^2} \mathbf{X}^{\mathsf{T}} \mathbf{X} + \mathbf{S}^{-1} \right] \mathbf{w}$$
$$\widehat{\mathbf{\Sigma}} = \left( \frac{1}{\sigma^2} \mathbf{X}^{\mathsf{T}} \mathbf{X} + \mathbf{S}^{-1} \right)^{-1}$$

Mean:

$$2\mathbf{w}^{\mathsf{T}}\mathbf{\Sigma}^{-1}\boldsymbol{\mu} = \frac{2}{\sigma^{2}}\mathbf{w}^{\mathsf{T}}\mathbf{X}^{\mathsf{T}}\mathbf{t}$$
$$\widehat{\boldsymbol{\mu}} = \frac{1}{\sigma^{2}}\widehat{\mathbf{\Sigma}}\mathbf{X}^{\mathsf{T}}\mathbf{t}$$

## Olympic example

► To make numbers better, rescape olympic year:

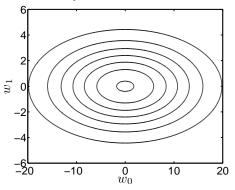
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# Olympic example

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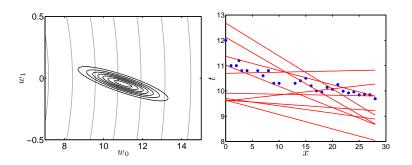
$$ightharpoonup$$
 1896 = 1, 1900 = 2, ..., 2008 = 27, 2012 = 28

▶ Prior density:



- ► Mean (0) and covariance (S).
- Quite a vague prior.

# Olympic example



Posterior (left) (prior shown in grey, zoomed in) and functions corresponding to some **w** sampled from posterior (right).

 Our motivation for being Bayesian was to be able to average predictions (at the test data x<sub>new</sub>) over all w

$$\mathbf{E}_{p(\mathbf{w}|\mathbf{X},\mathbf{t},\sigma^2)}\left\{f(\mathbf{w})\right\} = \int f(\mathbf{w})p(\mathbf{w}|\mathbf{t},\mathbf{X},\sigma^2) \ d\mathbf{w}$$

▶ We have the full posterior distribution over all possible values of w, it is also Gaussian and we computed the parameters.

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- ▶ We have the full posterior distribution over all possible values of w, it is also Gaussian and we computed the parameters.
- We can compute exactly the predictive density to make probabilistic predictions:

$$\begin{split} \rho(t_{\text{new}}|\mathbf{X},\mathbf{t},\mathbf{x}_{\text{new}},\sigma^2) &= \mathbf{E}_{p(\mathbf{w}|\mathbf{X},\mathbf{t},\sigma^2)} \left\{ p(t_{\text{new}}|\mathbf{x}_{\text{new}},\mathbf{w},\sigma^2) \right\} \\ &= \int p(t_{\text{new}}|\mathbf{x}_{\text{new}},\mathbf{w},\sigma^2) p(\mathbf{w}|\mathbf{t},\mathbf{X},\sigma^2) \ d\mathbf{w} \end{split}$$

We can even compute exactly, the predictive density to make probabilistic predictions:

$$\begin{split} \rho(t_{\mathsf{new}}|\mathbf{X},\mathbf{t},\mathbf{x}_{\mathsf{new}},\sigma^2) &=& \mathbf{E}_{\rho(\mathbf{w}|\mathbf{X},\mathbf{t},\sigma^2)} \left\{ p(t_{\mathsf{new}}|\mathbf{x}_{\mathsf{new}},\mathbf{w},\sigma^2) \right\} \\ &=& \int p(t_{\mathsf{new}}|\mathbf{x}_{\mathsf{new}},\mathbf{w},\sigma^2) p(\mathbf{w}|\mathbf{t},\mathbf{X},\sigma^2) \ d\mathbf{w} \end{split}$$

 $p(t_{\text{new}}|\mathbf{x}_{\text{new}},\mathbf{w},\sigma^2)$  is defined by our model as the product of  $\mathbf{x}_{\text{new}}$  and  $\mathbf{w}$  with some additive Gaussian noise.

$$p(t_{\text{new}}|\mathbf{x}_{\text{new}}, \mathbf{w}, \sigma^2) = \mathcal{N}(\mathbf{x}_{\text{new}}^\mathsf{T} \mathbf{w}, \sigma^2)$$

▶ Because this expression and the posterior are both Gaussian, the result of expectation is another Gaussian.

$$p(t_{\mathsf{new}}|\mathbf{X}, \mathbf{t}, \mathbf{x}_{\mathsf{new}}, \sigma^2) = \mathcal{N}(\mathbf{x}_{\mathsf{new}}^\mathsf{T} \widehat{\boldsymbol{\mu}}, \ \sigma^2 + \mathbf{x}_{\mathsf{new}}^\mathsf{T} \widehat{\boldsymbol{\Sigma}} \mathbf{x}_{\mathsf{new}})$$



► Therefore, the predictive density is

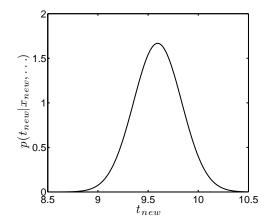
$$\textit{p}(\textit{t}_{\mathsf{new}}|\mathbf{X},\mathbf{t},\mathbf{x}_{\mathsf{new}},\sigma^2) = \mathcal{N}(\mathbf{x}_{\mathsf{new}}^{\mathsf{T}}\widehat{\boldsymbol{\mu}},\ \sigma^2 + \mathbf{x}_{\mathsf{new}}^{\mathsf{T}}\widehat{\boldsymbol{\Sigma}}\mathbf{x}_{\mathsf{new}})$$

where,

$$\widehat{oldsymbol{\Sigma}} = \left(rac{1}{\sigma^2} oldsymbol{\mathsf{X}}^\mathsf{T} oldsymbol{\mathsf{X}} + oldsymbol{\mathsf{S}}^{-1}
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and

$$\widehat{\boldsymbol{\mu}} = \frac{1}{\sigma^2} \widehat{\boldsymbol{\Sigma}} \mathbf{X}^\mathsf{T} \mathbf{t}.$$



Predictive density at 2012 Olympics. Note that  $\sigma^2$  was fixed at 0.05.

$$p(t_{\text{new}}|\mathbf{X}, \mathbf{t}, \mathbf{x}_{\text{new}}, \sigma^2) = \mathcal{N}(9.5951, 0.0572)$$

## Computing posterior: recipe

- (Assuming prior conjugate to likelihood)
- ▶ Write down prior times likelihood (ignoring any constant terms, i.e., the term that are irrelevant to w)
- Write down posterior (ignoring any constant terms)
- Re-arrange them so the look like one another
- Equate terms on both sides to read off parameter values.

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  - Prior effect will diminish as more data arrive.
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  - ► If we know nothing, can use a broad prior e.g. uniform density.

### Summary

- Moved away from a single parameter value.
- Saw how predictions could be made by averaging over all possible parameter values – Bayesian.
- Saw how Bayes rule allows us to get a density for w conditioned on the data (and other stuff).
- Computing the posterior is hard except in some cases....
- ....we can do it when things are conjugate.

### Recipe for a Bayesian linear model

- In the Bayesian linear regression, we compute a distribution over  $\mathbf{w}$  instead of estimating it by  $\widehat{\mathbf{w}} = (\mathbf{X}^\mathsf{T}\mathbf{X})^{-1}\mathbf{X}^\mathsf{T}\mathbf{t}$ .
- ► The model is

$$p(\mathbf{w}|\mathbf{X}, \mathbf{t}, \sigma^2) = \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}).$$

We use the Gaussian prior  $p(\mathbf{w})$  and the likelihood  $p(\mathbf{t}|\mathbf{w}, \mathbf{X}, \sigma^2) = \mathcal{N}(\mathbf{X}\mathbf{w}, \sigma^2\mathbf{I})$  to compute the model parameters  $\boldsymbol{\mu}$  and  $\boldsymbol{\Sigma}$ .

$$\widehat{oldsymbol{\Sigma}} = \left(rac{1}{\sigma^2} oldsymbol{\mathsf{X}}^\mathsf{T} oldsymbol{\mathsf{X}} + oldsymbol{\mathsf{S}}^{-1}
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$$\widehat{\mu} = \frac{1}{\sigma^2} \widehat{\mathbf{\Sigma}} \mathbf{X}^\mathsf{T} \mathbf{t}.$$

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Prediction (probabilistic predictions)

$$p(t_{\mathsf{new}}|\mathbf{X}, \mathbf{t}, \mathbf{x}_{\mathsf{new}}, \sigma^2) = \mathcal{N}(\mathbf{x}_{\mathsf{new}}^\mathsf{T} \widehat{\boldsymbol{\mu}}, \ \sigma^2 + \mathbf{x}_{\mathsf{new}}^\mathsf{T} \widehat{\boldsymbol{\Sigma}} \mathbf{x}_{\mathsf{new}})$$

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